



Kent Academic Repository

Key, Alastair, Roberts, David and Jari, Ivan (2021) *Reconstructing the full temporal range of archaeological phenomena from sparse data*. Journal of Archaeological Science, 135 . ISSN 0305-4403.

Downloaded from

<https://kar.kent.ac.uk/90224/> The University of Kent's Academic Repository KAR

The version of record is available from

<https://doi.org/10.1016/j.jas.2021.105479>

This document version

Author's Accepted Manuscript

DOI for this version

Licence for this version

CC BY-NC-ND (Attribution-NonCommercial-NoDerivatives)

Additional information

Versions of research works

Versions of Record

If this version is the version of record, it is the same as the published version available on the publisher's web site. Cite as the published version.

Author Accepted Manuscripts

If this document is identified as the Author Accepted Manuscript it is the version after peer review but before type setting, copy editing or publisher branding. Cite as Surname, Initial. (Year) 'Title of article'. To be published in *Title of Journal*, Volume and issue numbers [peer-reviewed accepted version]. Available at: DOI or URL (Accessed: date).

Enquiries

If you have questions about this document contact ResearchSupport@kent.ac.uk. Please include the URL of the record in KAR. If you believe that your, or a third party's rights have been compromised through this document please see our [Take Down policy](https://www.kent.ac.uk/guides/kar-the-kent-academic-repository#policies) (available from <https://www.kent.ac.uk/guides/kar-the-kent-academic-repository#policies>).

1
2
3
4
5
6
7
8
9
10
11
12
13
14
15
16
17
18
19
20
21
22
23
24
25
26
27
28
29
30
31
32
33
34
35
36
37
38
39

Reconstructing the full temporal range of archaeological phenomena from sparse data

Alastair Key ¹, David Roberts ², and Ivan Jarić ^{3,4},

Corresponding author: AK2389@cam.ac.uk

¹ Department of Archaeology, University of Cambridge, Cambridge, CB2 3DZ, UK

² Durrell Institute of Conservation and Ecology, School of Anthropology and Conservation, University of Kent, Canterbury, Kent, CT2 7NR, UK

³ Biology Centre of the Czech Academy of Sciences, Institute of Hydrobiology, České Budějovice, Czech Republic

⁴ University of South Bohemia, Faculty of Science, Department of Ecosystem Biology, České Budějovice, Czech Republic

Keywords:

Temporal distribution; modelling; end date and origin inference; optimal linear estimation; Weibull; computational archaeology; archaeological record

Abstract

Archaeologists rarely discover the first or last known occurrences of past cultural phenomena, yet ‘start’ and ‘end’ dates are central to our understanding of past human behaviour; therein lays a paradox long known within the discipline. Optimal linear estimation (OLE) has recently been used to reconstruct the full temporal range of prehistoric archaeological technologies using only the partial records available. That is, OLE has been used to reconstruct the portions of the archaeological record not yet evidenced through artefact discoveries. Here we present OLE to a wider archaeological audience and outline for the first time the model’s assumptions as they pertain to archaeological phenomena. We demonstrate OLE to be an accessible, user-friendly and methodologically transparent temporal range estimation method applied via a single set of equations. Further, we present five additional frequentist techniques that enable archaeologists to account for observation reliability, search effort and extreme data scarcity when inferring temporal ranges. These methods allow archaeologists to gain a more accurate understanding of the temporal range of past human behaviour.

1. Introduction

Recently introduced to archaeology from conservation science and palaeontological studies (e.g., Roberts and Solow, 2003; Pimiento and Clements, 2014; Zhang et al., 2020), optimal linear estimation (OLE) modelling is proving a useful technique to infer the full chronology of archaeological phenomena (Key et al, 2021a, 2021b; Bebber and Key, 2021). This brief communication is designed to introduce the technique to a wider archaeological audience, and to discuss how the model's predictions and assumptions interact with the unique nature of archaeological phenomena. We also present several additional, related techniques that are new to archaeological science and can be used separately or in conjunction with OLE to increase the accuracy of temporal estimates.

Unlike traditional assessments which use dated artefacts as a start or end point, OLE can infer how much longer a phenomenon is likely to have persisted prior to, or after, these known dates. That is, OLE is able to reconstruct the portion of the archaeological record that has not yet been discovered and provide a more accurate account of an archaeological phenomena's temporal presence (Fig. 1). To date, the technique has been used to extend the Oldowan and Acheulean periods by tens of thousands of years (Key et al., 2021a 2021b), and push back the origin of North American copper use by several hundred years (Bebber and Key, 2021).

The importance of reconstructing the earliest and latest portions of an archaeological phenomenon's temporal range has long been recognised (e.g., Cowgill, 1972; Surovell and Brantingham, 2007; Crema, 2012). Currently, Bayesian modelling approaches are the dominant (c.f. Buck and Meson, 2015; Ramsey, 2015) method applied to solving temporal probability questions in archaeological research. Although techniques addressing temporal uncertainty in archaeological occurrences are also well known (Crema, 2012; Kolář et al., 2016; Baxter and Cool, 2016). Most often used to calibrate radiocarbon dates (Bayliss, 2015; Crema and Bevan, 2021), Bayesian models can provide 'start' or 'end' dates for a given phenomenon based on probability densities derived from groups of temporally bounded radiocarbon samples (e.g., Wicks et al., 2014; Bicho et al., 2015). In turn, temporal range estimations have generally been limited to archaeological phenomena with associated radiocarbon dates.

Importantly, most Bayesian techniques applied to radiocarbon probability distributions do not directly model the start or end date of a phenomena. Instead, they provide revised probability estimates reconstructed from the probability densities of known radiocarbon samples; they do not directly reconstruct yet-to-be-discovered portions of the archaeological record. In other words, Bayesian techniques estimate when known artefact records occurred. Further, Bayesian techniques do not always consider changes to artefact occurrence frequencies through time (although it is possible [e.g., Ramsey, 2015; Fernández-López de Pablo and Barton, 2015; Banks et al., 2019; Crema and Kobayashi, 2020]).) Finally, Bayesian techniques (see: Otarola-Castillo and Torquato [2018] and references therein) are underpinned by the necessity of including prior assumptions about the phenomena they are investigating (Litton and Buck, 1995), meaning that inherent to any Bayesian temporal range estimates are subjective interpretations of prior information specified by the archaeologist (Pettitt and Zilhão, 2015) (although techniques to limit the impact of unverified *a priori* information can be applied [e.g., Long and Taylor, 2015]). While useful in many contexts, Bayesian techniques require prior information that can often be unavailable, are mathematically demanding, and can be difficult to compare to more traditional frequentist methods (Pettitt and Zilhão, 2015; Brook et al., 2019). It is for these reasons that their use has been limited within conservation studies (Boakes et al., 2015).

In contrast, OLE makes very few prior assumptions, and as a frequentist method based on the extreme value theory its predictions are rooted in the temporal spacing of the data that it investigates. This means that partial temporal records (i.e., periodic occurrences of artefacts through time) are not only easily accommodated into the method but are intrinsic to its predictions. Moreover, it can be applied to any type of dating information, any

89 archaeological phenomena, and at any timescale, so long as it is represented by lineal occurrences through time,
90 be this highly sporadic early stone tool technologies dated through optically stimulated luminescence (OSL),
91 faunal, $40\text{Ar}/39\text{Ar}$ or radiocarbon methods, or more recent archaeological phenomena dated through historical
92 texts or ethnographic sightings. Its implementation via a single set of equations further increases its transparency
93 and accessibility as an analytical tool. In other words, OLE has potential to be applied within any archaeological
94 context and by any archaeologist.

95
96 Other temporal range estimation techniques have been used in conservation science and paleobiology
97 research, and could potentially be used to reconstruct the archaeological record. However, some have been
98 shown to be outperformed by OLE (Rivadeneira et al., 2009) or require records to represent stationary Poisson
99 processes (i.e., there is no decline in frequency towards extinction, or increase in frequency after invention) (e.g.,
100 Solow, 1993), which is not realistic for most archaeological phenomena (Surovell and Brantingham, 2007;
101 Mesoudi and Lycett, 2009; Lycett, 2015; Mesoudi, 2015; Shennan, 2015). Indeed, one of the reasons that OLE
102 can be readily applied within archaeology is that, as with species, the origination or extinction of culture is not
103 sudden. It develops over time, meaning the likelihood of identifying sites that represent the start and end points
104 of a phenomena are incredibly low (Surovell and Brantingham, 2007; Prasciunas and Surovell, 2015). Rather,
105 phenomena are likely to have persisted for longer periods, and these periods will be relatively difficult to detect
106 in the archaeological record. There are always exceptions, such as catastrophic events that quickly wipe out a
107 population, but these will be less common.

108
109 It is important to note that although first developed for conservation science (Roberts and Solow, 2003;
110 Solow, 2005), OLE has no parameters specific to biological organisms and can readily be applied to cultural
111 traditions. Moreover, the analogous mechanisms underpinning biological and cultural evolution allow for similar
112 factors to be influencing the start/end dates of both (Lycett, 2015; Mesoudi, 2015; Shennan, 2015). In addition
113 to OLE, we also briefly describe here five other techniques that enable researchers to account for observation
114 reliability (Jarić and Roberts, 2014; Brook et al., 2019), search effort (McCarthy, 1998), and situations with
115 extreme data scarcity, such as datasets consisted of only two records (Solow and Roberts, 2003) or just a single
116 record (Roberts and Jarić, 2020), when inferring temporal ranges.

117 118 2. How does optimal linear estimation (OLE) work?

119
120 OLE requires the oldest or youngest dated occurrences of a phenomena to be entered into the model
121 (depending on whether it is being used to estimate a 'start' or an 'end'), from which the timings and chronological
122 spacing of these known occurrences are used to statistically estimate how much earlier or longer the phenomena
123 is likely to have existed. Ten dates are generally recommended as being optimal (Solow, 2005; Rivadeneira et al.,
124 2009), although lower sample sizes (e.g., $n = 5$) have also been demonstrated to display good accuracy (Clements
125 et al., 2013). OLE relies on the assumption that the dates entered into the model display (at least roughly) a joint
126 distribution with a 'Weibull form'. As noted above, this is a valid assumption for most archaeological phenomena.
127 In turn, the shape (form) parameters of the Weibull distribution used in the OLE model are based on the
128 chronology (spacing) of the dates entered. An 'end' or 'start' point can then be determined and is defined as the
129 point at which the Weibull distribution determines that another occurrence should have been found had the
130 cultural phenomenon not ended or not yet existed (relative to the temporal direction of the model) (Fig. 1).

131
132 In case of a phenomenon ending, the model predicts that given the temporal spacing of the known
133 archaeological record, we would have expected to have found a younger archaeological site relative to the
134 youngest currently known, if the phenomenon still existed beyond the tail end of the model's distribution. Given
135 that no sites have been found, we can infer the phenomenon ends at this point. In cases of origination, the model
136 infers that given the temporal spacing of the known archaeological record, another earlier occurrence of the

137 phenomenon would have been expected to have been found, relative to the currently known oldest occurrence,
138 if it existed beyond the tail of the modelled distribution.

139

140 2.1 Assumptions

141

142 Although OLE has few underlying assumptions compared to other temporal modelling techniques (Solow,
143 2005; Rivadeneira et al., 2009; Clements et al., 2013; Pettitt and Zilhão, 2015; Otarola-Castillo and Torquato,
144 2018), there are several to be aware of. First, OLE assumes a continuation of the phenomena in question after or
145 before the latest or earliest (respectively) currently known occurrences. Simply put, it assumes that it is unlikely
146 that the most recent or earliest observation of a process would be the last/first point in time when the process
147 was active. This can be safely assumed for most archaeological phenomena (see above), but near instantaneous
148 ends to cultural phenomena are possible. This could range from catastrophic natural disasters (e.g., volcanic
149 eruptions) through to the rapid replacement of one artefact type with another (e.g., changes to coinage). In
150 instances where this is assumed, other methods, designed for cases with constant sighting rates prior to a
151 phenomenon's end, could be considered instead, such as that by Solow (1993). However, it should be noted that
152 OLE and other frequentist methods are sensitive to, and account for, the amount of available data on the studied
153 phenomenon. For example, in cases of well-studied and recent artefact types, where occurrences are temporally
154 densely distributed and frequent, predicted end/start points would shift to and closely match the time of the
155 last known record.

156

157 Second, OLE assumes all observations (in this case archaeological occurrences) to be discrete. When
158 modelling species extinction this means that each observation is assumed to represent a different, independent
159 sighting event. Given that artefacts represent physical manifestations of cultural information contained within a
160 biological repository, and ultimately it is the chronology of the cultural information that is being modelled, a
161 'discrete' archaeological occurrence is at its most extreme an artefact/assemblage assumed to be produced by a
162 unique individual relative to those already included in the model (Bebber and Key, 2021). This means that all
163 artefactual occurrences should be assumed to have been made by different individuals. At its most moderate,
164 'discrete' may simply refer to an independent representation of the phenomena; although this entails an
165 assumption of random sampling from a population. Several identical dates can be used so long as they all meet
166 this criterion. Indeed, variation in data type, quality and characteristics needs to be considered carefully to avoid
167 biases when applying OLE (as with any modelling technique). Highly clustered data make an important example,
168 and the archaeological record is full of waste middens, battle grounds, occupancy sites, artefact caches and other
169 phenomena that should be considered carefully in terms of what precisely they represent when included within
170 the model. In cases where cultural phenomena are broader and the product of multiple
171 individuals/craftsperson's, such as specific settlement types or some ceremonial outfits, then it may be more
172 appropriate to consider discrete observations at a population level.

173

174 Third, although OLE does not assume the probability of detection to be consistent, it does assume relatively
175 stable search effort that never equates to zero. When modelling species extinction, this means that any breaks
176 in between dates in the model (occurrences) are not a function of irregular search effort. In other words, the
177 model assumes that gaps in species sighting data are not because people are not looking for the species. This
178 assumption is less relevant to studies of paleontological or archaeological phenomena, as the temporal record
179 of fossils and artefacts through time (and therefore their spacing) are not dictated by human search effort, but
180 geological and taphonomic factors in combination with frequency changes (Surovell and Brantingham, 2007;
181 Surovell et al., 2009). Archaeologists can display biases in search effort towards sediments of a specific age, but
182 assuming that there is no intentional avoidance of artefacts then this assumption should not be violated in most
183 archaeological instances (notably, geographic biases in search effort may also influence temporal records).

184

185 Although OLE makes no other assumptions, Clements et al. (2013) highlighted two additional factors that
186 could potentially influence the model's accuracy. The first includes sudden changes to the rate of a phenomena's
187 decline/uptake. Within an archaeological context, this means the addition of new pressures could alter the rate
188 of cultural decline/uptake experienced for some, but not all, occurrences in the model. For example, sudden
189 environmental change relevant to only the latest occurrences of an 'extinction' (end) model could additionally
190 speed up the phenomena's decline due to additional stresses on the population. The second factor concerns
191 changes to the 'observability' of a phenomena. This does not relate to the naturally increasing/decreasing
192 likelihood of discovering an occurrence as the phenomena become more or less widespread (Surovell and
193 Brantingham, 2007; Mesoudi and Lycett, 2009; Jordan and Cummings, 2014). Instead, it refers to a change in
194 relative observability. For instance, taphonomic processes can introduce natural bias to the preservation of
195 archaeological sites, which can impact observability irrespective of the demographic and frequency trends
196 present at the time of existence (Surovell and Brantingham, 2007; Surovell et al., 2009). Further, spatial
197 considerations may also impact observability through changes to the accessibility of past landscapes (for
198 example, earlier or later occurrences may be more or less likely to be discovered due to flooding events [e.g., the
199 flooding of Doggerland]). Thus, these factors should be assumed constant when using OLE estimates (see Solow
200 [2005], Clements et al. [2013], and Boakes et al. [2015] for further details).

201
202 Finally, there are several expectations relevant to using OLE in archaeology that are not necessarily
203 applicable in other contexts. First, there is an inherent assumption that the dates used in the model are an
204 accurate representation of the temporal presence of the occurrence in question (see: Crema, 2012). That is, the
205 dates associated with an artefact/site should accurately reflect when it entered the archaeological record (or at
206 least, as accurate as the relevant dating method can be). Finally, there should be sufficient accuracy of an
207 archaeological occurrence's identification when including it in an OLE model. This is primarily an issue when
208 typological artefact classifications are subjective and have potential to vary between analysts; after all, one
209 cannot identify the final or first occurrence of a phenomena if it cannot be reliably detected. However, this is
210 likely more an issue for some archaeological fields than others.

211 212 2.2 OLE model

213
214 The formulaic expression of OLE has recently been published in open access archaeological literature (Key
215 et al., 2021a, 2021b), and is widely available elsewhere, including the original articles describing the technique
216 (Roberts and Solow, 2003; Solow, 2005). Further, the R sExtinct software package provides an easily accessible
217 means through which to run OLE (Clements, 2013). We provide an explanation of the OLE equations in the
218 attached Supplementary Information, along with a link to the R script provided by Clements (2013). Notably,
219 models inferring the 'start' date of archaeological phenomena need to be adjusted to run in the reverse temporal
220 direction to those provided by Clements (2013; cf. Key et al., 2021b). The 10 youngest or oldest dates should be
221 approximately used as the beginning of the period, dependent on the direction of the model. As with any
222 frequentist model the most common time unit used is years, but decades, centuries or millennia can be used
223 just as well.

224
225 As noted above, OLE provides temporal estimates based on extreme Weibull form distributions that are
226 tailored relative to the dated occurrences used, their distribution, and intervals observed between these dates.
227 The distribution's curve is then used to identify an 'end point' or 'point of origin' beyond which the phenomenon
228 can be inferred to no longer exist, as if it did (given current artefact distributions) another artefactual occurrence
229 would be known about between the last/earliest current known occurrence and the inferred point of end/origin.
230 To put it another way, OLE asks, given the distribution of known occurrences, how likely is it that another does
231 not exist in the archaeological record.

232

233 Two estimated dates are commonly produced by OLE. One represents the estimated origin (T_O) or end (T_E)
234 date of the phenomenon in question. The other represents the upper bound of each model's confidence interval
235 (T_{CI}), although lower bounds can also easily be produced and applied (Roberts and Solow, 2003). T_O and T_E dates
236 are the main output of archaeological OLE models and represent the 'end' or 'origin' point in whatever dating
237 time scale used for the input data (e.g., BP, BCE). T_{CI} dates represent the point beyond which there is a 5% or less
238 probability (in line with $\alpha = 0.05$) that the phenomenon still existed. In other words, T_{CI} dates represent a
239 confidence limit where in 95 out of 100 cases, the true end/origin point will be within the confidence interval.

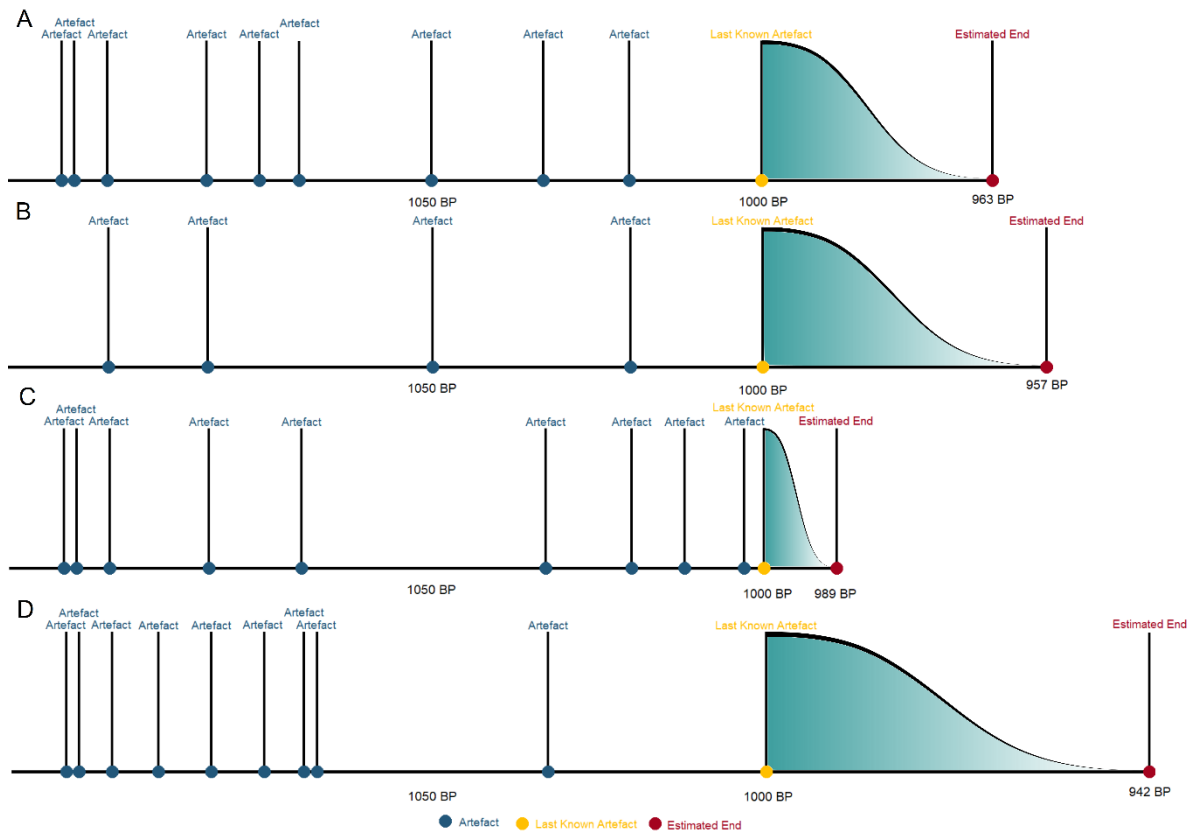
240
241 Archaeological dating methods often come with a degree of uncertainty. That is, although the age of
242 artefactual occurrences can be identified, they often come with error ranges, or date ranges with varying
243 distributions of likelihood. Combined with the OLE method's use of discrete time units (e.g., specific years), it
244 means that point estimates must be identified and used in place of these ranges. It is possible to run individual
245 OLE models using central tendency values taken from ranges, and this will provide a reasonable estimated
246 account of the phenomenon's temporal range (particularly if the date range has an approximate normal
247 distribution of likelihood). Other methods for deciding on point estimates can be used, so long as they are
248 justified (e.g., a secondary dating technique indicating that a specific portion of the first techniques range is more
249 likely). In other instances, however, it may be favourable to account for temporal range uncertainty through
250 repeated sampling procedures. As already applied elsewhere (Bebber and Key, 2021; Key et al., 2021), it is
251 possible to draw dates randomly from a normal or uniform distribution (or other distribution, if preferred) within
252 a defined range. This is repeated for all investigated occurrences and the randomly generated datasets are
253 subsequently assessed with the OLE method. By repeating this procedure many times (e.g., 10,000) and using
254 averages from the OLE data produced, it becomes possible to account for the temporal uncertainty associated
255 with many archaeological dating techniques. Additional frequentist techniques that specifically account for
256 record reliability, including dating uncertainty, are described below (Section 3.1).

257 258 259 **2.3 How does temporal spacing of dated archaeological occurrences impact OLE estimations?**

260
261 OLE has repeatedly been demonstrated to be robust within a variety of scenarios, including those that vary
262 in temporal scale, 'sighting' probabilities, or search effort and trajectories (Rivadeneira et al., 2009; Clements et
263 al., 2013). This means that OLE tends to remain accurate under most archaeological scenarios. As with any
264 statistical modelling, however, the characteristics of the data entered into the model determine the output and
265 the inferences derived thereof. Thus, it is useful to highlight how variation in sampling patterns and discovery
266 happenstance can influence temporal range estimations. Figure 1 details a typical sequence of ten archaeological
267 finds through time, under four different scenarios. This is a simple example and does not represent a test of
268 robustness, as this has already been undertaken elsewhere (see Rivadeneira et al. 2009; Clements et al., 2013).
269 Nevertheless, it is a useful illustration of the method's performance, as it highlights its sensitivity to the
270 distribution and trends in records.

271
272 Figure 1a details a typical artefact sequence where there is a greater number of earlier artefactual
273 occurrences (blue), but as the phenomena increasingly moves towards its end the likelihood of finding
274 artefacts/sites decreases until there is a final known occurrence (yellow). Using this example, OLE models predict
275 the phenomena to end at 965 years before present (BP), 35 years after the youngest known artefact (1000 BP).
276 Using the same artefactual occurrences but with several being randomly removed (such that $n = 5$, while ensuring
277 the youngest date remains the same), the model's inferred end date shifts slightly forward to 957 BP (Fig. 1b),
278 suggesting a slightly longer persistence of the phenomenon. Of course, this is a simple example, but it highlights
279 the ability of the method to be used with low sample sizes, albeit with increased uncertainty.

281 The factor most likely to influence an OLE model's prediction is the temporal density of dates close to the
 282 latest or earliest known occurrence (depending on its temporal direction). If there are greater numbers of known
 283 occurrences immediately prior to the last one, then the model will predict a closer end date (989 BP; Fig. 1c).
 284 Indeed, it follows that if there are several closely dated sites and then an absence of occurrences, it is likely that
 285 the last known record will be reasonably close to the true end date of the phenomenon. Conversely, relatively
 286 few records prior to the last known occurrence will point to a likely declining trend of occurrence, and that the
 287 phenomenon likely continued for an extended period with low chance of detection. Thus, in these instances the
 288 model has increased uncertainty and predicts a relatively longer tail of presence for the phenomenon (942 BP;
 289 Fig. 1d).
 290



291
 292 **Figure 1:** Examples of how OLE temporal estimates are influenced by date distributions (these illustrations are
 293 derived from real OLE estimates). Figure A represents a 'standard' archaeological scenario where more
 294 occurrences exist closer to the phenomenon's peak and then gradually decrease prior to the 'last known artefact'
 295 ($n = 10$). Figure B represents the same scenario, but happenstance and/or reduced search effort has resulted in
 296 only five artefacts (and therefore dates) being available ($n = 5$; randomly sampled from Figure A's scenario).
 297 Figure C illustrates improved surveying and/or preservation conditions, which has resulted in a greater number
 298 of instances having been found prior to the last known artefact ($n = 10$). This will predict a faster end to the
 299 phenomenon; the model's logic is that *if* it had continued beyond 989 BP, archaeologists should have discovered
 300 another site between 1000 – 989 BP given the density and distribution of sites found prior to 1000 BP. Figure D
 301 demonstrates the reverse scenario, where fewer dates are available around the 'last known artefact', indicating
 302 a declining trend of occurrence. In such a scenario, a quick end to the phenomenon is unlikely due to extended
 303 intervals between occurrences, and therefore the increased likelihood that artefacts do exist but remain
 304 undiscovered. Note that scenarios C and D are relatively extreme examples for demonstrative purposes.
 305

306 3. Further frequentist range estimation methods

307

3.1 Methods that account for record reliability

Archaeological records often contain artefacts of variable reliability. This may stem from their state of preservation and other factors impeding identification, uncertainty of their origin (e.g., surface finds), or credibility of the person reporting the find. Whether an artefact or site is accepted as a valid record can strongly affect the inferences derived using temporal estimation methods (Roberts et al., 2010). Thus, it is important for the reliability of occurrence identification to be considered as part of some archaeological scenarios that apply these techniques.

The majority of the existing temporal estimation methods designed to handle mixed-certainty records either classify them as reliable or unreliable, or assign them a score that represents probability that a record is true, based on the record type and its characteristics (Boakes et al., 2015; Brook et al., 2019). While the common approach for treating variable reliability is to include or discard records based on their type or arbitrary inclusion probability thresholds, recently developed approaches by Jarić and Roberts (2014) and Brook et al. (2019) allow for individual record reliabilities to be directly included in existing temporal inference methods (including OLE).

The approach by Jarić and Roberts (2014) represents a simple modification of existing methods and allows each individual record to be weighted based on its reliability score. Although so far only applied to the standard Solow method (Solow, 1993), Jarić and Roberts (2014) suggest that it can also be readily applied to other frequentist methods, including those by Strauss and Sadler (1989), Solow and Roberts (2003) and McInerney et al. (2006). The modification works by replacing standard record time series, represented by a binary sequence of presences and absences, with presences expressed as probabilities that represent the reliability of each individual record (i.e., a likelihood that the given observation is true) (Jarić and Roberts, 2014). In effect, the number of records in a dataset is replaced by a sum of probability values assigned to each record, indicating the most likely number of observations. Additionally, individual record reliabilities are used to estimate the most likely endpoint of a sighting record, based on the likelihood that all later records in the dataset are false. Brook et al. (2019) introduced additional improvements to this approach by also estimating the likely year of the first true record, based on the likelihood that all preceding records are false. Together, these approaches can substantially improve the predictive power of origin and end-point (extinction) estimation methods when occurrence records display mixed-certainty.

A further simple and powerful extension to OLE and other existing temporal range estimation methods was proposed by Brook et al. (2019). The approach is based on resampling without replacement from a dataset of records, where reliability of each record is used as its probability of being sampled and included. Repeated sampling and extinction inference based on each sample record produces a range and frequency distribution around the extinction date. The method was demonstrated to be robust and to perform well, and has a number of advantages compared to existing methods: 1) it is based on resampling of records and these can thereafter be used as input for other end/origin estimation methods, including OLE; 2) it is simple and easily applicable using the freely available R script; 3) it allows direct integration of overlapping records (Brook et al., 2019).

3.2 Methods that account for collection effort

As discussed previously, relatively stable collection effort represents one of the key assumptions of OLE and other methods from this group. Any changes in collection effort over geological time will be confounded with true changes and trends in the presence and frequency of phenomenon studied, and might consequently lead to biased results. We note this to potentially be less of an issue for archaeology relative to other fields using these techniques (Section 2.1), but there are situations where collection efforts may be biased in favour of older or more recent sediments due to temporal or geographic biases.

357 When a substantial level of temporal instability in collection effort is suspected, methods such as that
358 introduced by McCarthy (1998) should be considered. Also known as the "Partial Solow equation", this method
359 represents a modification of the Solow method (Solow, 1993) and uses indices of collection effort made in each
360 discrete time unit (i.e., year). It can be used to estimate the likelihood that a phenomenon has ended, and the
361 likely extinction time, based on the proportion of collection effort made prior to the last known record and the
362 total collection effort, as well as the total number of records.

364 3.3 Methods designed for extreme data scarcity

365
366 Archaeological records can be scarce, which can complicate or obstruct application of quantitative methods
367 such as OLE. Overall, sample sizes below five records are not recommended for OLE and other standard methods
368 from this group, and usually cannot be applied to less than three records. However, two extinction inference
369 methods presented here are designed to handle such extreme data scarcity. A non-parametric method
370 introduced by Solow and Roberts (2003), and based on the truncation point estimate by Robson and Whitlock
371 (1964), provides inference of the extinction date based on the timing of the last two records. Here, extinction
372 likelihood simply represents a proportion of the interval between the last two records, and the interval between
373 the second-last record and the end of the observation period. This technique was recently applied to estimate
374 the Lomekwian's date of origin (Key et al., 2021b).

375
376 Roberts and Jarić (2020) further proposed an approach based on the Partial Solow equation (McCarthy, 1998)
377 that is able to infer extinction probability and confidence intervals for phenomena known from only a single
378 record. Using indices of collection effort, extinction likelihood represents a proportion between the collection
379 effort made prior to the only known record and the total effort over the whole collection period (Roberts and
380 Jarić, 2020).

381
382 Considering the scarcity of input data, both methods are very conservative, and produce wide confidence
383 intervals. Nevertheless, for such data-deficient situations, these methods often represent the only available
384 quantitative indices for the temporal spans of phenomena, and can be used as just one line of evidence, or as a
385 preliminary screening tool (Roberts and Jarić, 2020). New archaeological phenomena or exceptional finds very
386 rarely preserved (e.g., Lower and Middle Palaeolithic organic technology) would benefit most from these
387 methods.

388 4. Discussion and Conclusion

389
390
391 The adage 'garbage in, garbage out' applies equally to optimal linear estimation as it does to any other modelling
392 technique. Following Pettitt and Zilhão's (2015) example, we wanted to emphasise the assumptions underlying
393 the models detailed here and used elsewhere (Key et al., 2021a, 2021b), and to suggest points of good practice
394 for those wishing to use them. There are no 'one rule fits all' approaches to meeting all assumptions, and in our
395 own experience we have taken the decision to slightly modify how we define discrete archaeological occurrences
396 depending on the context (e.g., Bebbber and Key, 2021). However, by detailing these suggestions as the methods
397 are introduced, we hope to create an environment where OLE can be applied to the archaeological record with
398 minimal room for misleading and inaccurate results.

399
400 Predictive models are, of course, just that; predictions. They can be based on the best available evidence, but it
401 is nevertheless the case that unexpected or unknowable influences could have been acting on the archaeological
402 phenomenon in question, or that key information has been obscured to date due to human-based search
403 biases/effort or taphonomic processes. Future discoveries therefore have the potential to overhaul any
404 predictions made using the techniques outlined here by extending temporal ranges or increasing/decreasing
405 predicted slopes of decline/growth. This should not be viewed as a failure on a part of the model. Instead, it is

406 an invitation to reapply the model using these new data, and to continuously update estimated temporal ranges
407 as new information comes to light.

408

409 OLE is an accessible, user-friendly and methodologically transparent temporal range estimation method applied
410 via a single set of equations. It is for these reasons it has become widely used within ecological and
411 palaeontological studies (Rout et al., 2010; Lee, 2014). We therefore hope that OLE and the other techniques
412 described here become accepted and applied within archaeological research. Doing so will allow us to gain a
413 better understanding the temporal presence of past human behaviour.

414

415

416 **Acknowledgments**

417

418 A.K.'s research is supported by the Royal Society. I.J. was supported by J.E. Purkyně Fellowship of the Czech
419 Academy of Sciences.

420

421 **References**

422

423 Banks, W.E., Bertran, P., Ducasse, S., Klaric, L., Lanos, P., Renard, C., Mesa, M., 2019. An application of
424 hierarchical Bayesian modeling to better constrain the chronologies of Upper Paleolithic archaeological cultures
425 in France between ca. 32,000–21,000 calibrated years before present. *Quat. Sci. Rev.* 220, 118–214

426

427 Baxter, M.J. and Cool, H.E.M., 2016. Reinventing the wheel? Modelling temporal uncertainty with applications
428 to brooch distributions in Roman Britain. *J. Arch. Sci.* 66, 120-127

429

430 Bayliss, A., 2015. Quality in Bayesian chronological models in archaeology. *World Arch.* 47 (4), 677–700

431

432 Bebber, M.R., Key, A.J.M., in press. Optimal linear estimation (OLE) modelling supports early Holocene (9000-
433 8000 RCYBP) copper tool production in North America. *Am. Antiq.* DOI: TBC

434

435 Bicho, N., Marreiros, J., Cascalheira, J., Pereira, T., Haws, J., 2015. Bayesian modeling and the chronology of the
436 Portuguese Gravettian. *Quat. Intern.* 359-360: 499–509

437

438 Buck, C.E. and Meson, B. 2015. On being a good Bayesian. *World Arch.* 47 (4), 567–584

439

440 Boakes, E.H., Rout, T.M., Collen, B., 2015. Inferring species extinction: the use of sighting records. *Methods in*
441 *Ecol. Evo.* 6 (6), 678–687

442

443 Brook, B.W., Buettel, J.C., Jarić, I., 2019. A fast re-sampling method for using reliability ratings of sightings with
444 extinction-date estimators. *Ecology* 100 (9), e02787

445

446 Buck, C.E., Meson, B., 2015. On being a *good* Bayesian. *World Arch.* 47 (4), 567–584

447

448 Clements C., 2013. sExtinct: Calculates the historic date of extinction given a series of sighting events. R
449 package version 1.1. <https://cran.r-project.org/src/contrib/Archive/sExtinct/>, Accessed 16th Jul 2020

450

451 Clements, C.F., Worsfold, N.T., Warren, P.H., Collen, B., Clark, N., Blackburn, T.M., Petchey, O.L., 2013.
452 Experimentally testing the accuracy of an extinction estimator: Solow's optimal linear estimation model. *J.*
453 *Animal Ecol.* 82 (2), 345–354

454

455 Cowgill, G.L., 1972. Models, methods and techniques for seriation. In: Clarke, D.L. (Ed.) *Models in Archaeology.*
456 Methuen and Co., London. pp. 381–424

457

458 Crema, E.R. 2012. Modelling temporal uncertainty in archaeological analysis. *J. Arch. Meth. Theory* 19, 440-461

459

460 Crema, E.R., Kobayashi, K. 2020. A multi-proxy inference of Jōmon population dynamics using bayesian phase
461 models, residential data, and summed probability distribution of 14C dates. *J. Arch. Sci.* 117, 105136
462
463 Crema, E.R., Bevan, A. 2021. Inference from large sets of radiocarbon dates: software and methods.
464 *Radiocarbon* 61 (1), 23-39
465
466 Fernández-López de Pablo, J., Barton, C.M., 2015. Bayesian estimation dating of lithic surface collections. *J.*
467 *Arch. Meth. Theory* 22, 559–583
468
469 Jarić, I., Roberts, D.L., 2014. Accounting for observation reliability when inferring extinction based on sighting
470 records. *Biodivers. Conserv.* 23 (11), 2801–2815
471
472 Jordan, P., Cummings, V., 2014. Prehistoric hunter-gatherer innovations. In: Cummings, V., Jordan, P., Zvelebil,
473 M. (Eds.) *The Oxford Handbook of the Archaeology and Anthropology of Hunter-Gatherers*. Oxford University
474 Press, Oxford. pp. 585–606
475
476 Key, A.J.M., Jarić, I., Roberts, D.L., 2021a. Modelling the end of the Acheulean at global and continental levels
477 suggests widespread persistence into the Middle Palaeolithic. *Humanities Soc. Sci. Comm.* 8, 55
478
479 Key, A.J.M, Roberts, D.L, Jarić, I., 2021b. Statistical inference of earlier origins for the first flaked stone
480 technologies. *J. Human Evo.* 154, 102976
481
482 Kolář, J., Macek, M., Tkáč, P., Szabó, P. 2016. Spatio-temporal modelling as a way to reconstruct patterns of past
483 human activities. *Archaeometry* 58 (3), 513-528
484
485 Lee, T.E., McCarthy, M.A., Wintle, B.A., Bode, M., Roberts, D.L., Burgman, M.A., 2014. Inferring extinctions from
486 sighting records of variable reliability. *J. Applied Ecol.* 51 (1), 251–258
487
488 Litton, C.D., Buck, C.E., 1995. The Bayesian approach to the interpretation of archaeological data. *Archaeom.* 37
489 (1), 1–24
490
491 Long, T., Taylor, D., 2015. A revised chronology for the archaeology of the lower Yangtze, China, based on
492 Bayesian statistical modelling. *J. Arch. Sci.* 63, 115–121
493
494 Lycett, S.J., 2015. Cultural evolutionary approaches to artifact variation over time and space: basis, progress,
495 and prospects. *J. Arch. Sci.* 56, 21–31
496
497 McCarthy, M.A., 1998. Identifying declining and threatened species with museum data. *Bio. Conserv.* 83 (1), 9–
498 17
499
500 McInerney, G.J., Roberts, D.L., Davy, A.J., Cribb, P.J., 2006. Significance of sighting rate in inferring extinction and
501 threat. *Conserv. Bio.* 20 (2), 562–567
502
503 Mesoudi, A., 2015. Cultural evolution: a review of theory, findings and controversies. *Evo. Biology* 43, 481–497
504
505 Mesoudi, A., Lycett, S.J., 2009. Random copying, frequency-dependent copying and culture change. *Evo.*
506 *Human Behav.* 30 (1), 41–48
507
508 Otarola-Castillo, E., Torquato, M.G., 2018. Bayesian statistics in archaeology. *Annual R. Anth.* 47, 435–453
509
510 Pettitt, P., Zilhão, J., 2015. Problematizing Bayesian approaches to prehistoric chronologies. *World Arch.* 47 (4),
511 525–542
512
513 Pimiento, C., Clements, C.F., 2014. When Did *Carcharocles megalodon* Become Extinct? A New Analysis of the
514 Fossil Record. *PLOS One* 9 (1), e111086
515

516 Prasciunas, M.M., Surovell, T.A., 2015. Reevaluating the duration of Clovis: The problem of non-representative
517 radiocarbon. In: Smallwood, A.M. and Jennings, T.A. (Eds.) *Clovis: On the Edge of a New Understanding*. Texas
518 A&M University Press, College Station. pp. 21–35.

519

520 Ramsey, C.B., 2015. Bayesian approaches to the building of archaeological chronologies. In: Barcelo, J.A. and
521 Bogdanovic, I. (Eds.) *Mathematics and Archaeology*. CRC Press, Boca Raton. pp. 272–292

522

523 Rivadeneira, M.M., Hunt, G., Roy, K., 2009. The use of sighting records to infer species extinctions: an
524 evaluation of different methods. *Ecology* 90 (5), 1291–1300

525

526 Roberts, D.L., Solow, A.R., 2003. When did the dodo become extinct? *Nature* 426, 245

527

528 Roberts, D.L., Elphick, C.S., Reed, J.M., 2010. Identifying anomalous reports of putatively extinct species and
529 why it matters. *Conserv. Bio.* 24 (1), 189–196

530

531 Roberts, D.L., Jarić, I., 2020. Inferring the extinction of species known only from a single specimen. *Oryx* 54 (2),
532 161–166

533

534 Robson, D.S., Whitlock, J.H., 1964. Estimation of a truncation point. *Biometrika* 51 (1/2), 33–39

535

536 Rout, T.M., Heinze, D., McCarthy, M.A., 2010. Optimal allocation of conservation resources to species that may
537 be extinct. *Conserv. Bio.* 24, 1111–1118

538

539 Shennan, S., 2015. Demography and cultural evolution. In: Scott, R.A. and Buchmann, M.C. (Eds.) *Emerging*
540 *Trends in the Social and Behavioral Sciences*. John Wiley and Sons, Hoboken. pp. 1–14

541

542 Solow, A.R., 1993. Inferring extinction from sighting data. *Ecology* 74 (3), 962–964

543

544 Solow, A.R., 2005. Inferring extinction from a sighting record. *Math. Biosciences* 195 (1), 47–55

545

546 Solow, A.R., Roberts, D.L., 2003. A nonparametric test for extinction based on a sighting record. *Ecology* 84 (5),
547 1329–1332

548

549 Strauss, D., Sadler, P.M., 1989. Classical confidence intervals and Bayesian probability estimates for ends of
550 local taxon ranges. *Math. Geology* 21 (4), 411–427

551

552 Surovell, T.A., Brantingham, P.J., 2007. A note on the use of temporal frequency distribution in studies of
553 prehistoric demography. *J. Arch. Sci.* 34 (11), 1868–1877

554

555 Surovell, T.A., Finley, J.B., Smith, G.M., Brantingham, P.J., Kelly, R., 2009. Correcting temporal frequency
556 distribution for taphonomic bias. *J. Arch. Sci.* 36 (8), 1715–1724

557

558 Wicks, K., Pirie, A., Mithen, S.J., 2014. Settlement patterns in the late Mesolithic of western Scotland: the
559 implications of Bayesian analysis of radiocarbon dates and inter-site technological comparisons. *J. Arch. Sci.* 41,
560 406–422

561

562 Wurzer, G., Kowarik, K., Reschreiter, H., 2015. *Agent-Based Modeling and Simulation in Archaeology*. Springer,
563 Cham

564

565 Zhang, H., Jarić, I., Roberts, D.L., He, Y., Du, H., Wu, J., Wang, C., Wei, Q., 2020. Extinction of one of the world's
566 largest freshwater fishes: Lessons for conserving the endangered Yangtze fauna. *Sci. Total Environ.* 710, 136242