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DOI

<https://doi.org/10.7554/eLife.33468>

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Large-scale replication study reveals a limit on probabilistic prediction in language comprehension

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ABSTRACT

In current theories of language comprehension, people routinely and implicitly predict upcoming words by pre-activating their meaning, morpho-syntactic features and even their specific phonological form. To date the strongest evidence for phonological prediction comes from a landmark 2005 Nature Neuroscience publication by DeLong, Urbach and Kutas, who observed a graded modulation of electrical brain potentials (N400) to nouns and preceding articles by the probability that people use a word to continue the sentence fragment ('cloze'). In a direct replication study spanning 9 laboratories (N=334), we failed to replicate the crucial article-elicited N400 modulation by cloze, while we successfully replicated the commonly-reported noun-elicited N400 modulation. We observed this pattern of failure and success in a pre-registered replication analysis, a pre-registered single-trial analysis, and exploratory Bayesian analyses. Contra the strong prediction view in which people routinely pre-activate the phonological word-form, our results suggest a more limited role for prediction during language comprehension.

1 INTRODUCTION

2 In the last decades, the idea that people routinely and implicitly predict upcoming
3 words during language comprehension turned from a highly controversial hypothesis to a
4 widely accepted assumption. Initial objections to prediction in language were based on a lack
5 of empirical support (e.g., Zwitserlood, 1989), incompatibility with traditional bottom-up
6 models and contemporary interactive models of language comprehension (e.g., Kintsch,
7 1988; Marslen-Wilson & Tyler, 1988), and the purported futility of prediction in a generative
8 system where sentences can continue in infinitely many different ways (Jackendoff, 2002).
9 Current theories of language comprehension, however, reject such objections and posit
10 prediction as an integral and inevitable mechanism by which comprehension proceeds
11 quickly and incrementally (e.g., Altmann & Mirkovic, 2009; Dell & Chang, 2014; Pickering
12 & Garrod, 2013). Prediction, i.e., context-based pre-activation of an upcoming linguistic
13 input, is thought to occur at all levels of linguistic representation (semantic, morpho-syntactic
14 and phonological/orthographic) and serves to facilitate the integration of newly available
15 bottom-up information into the unfolding sentence- or discourse-representation. In this line of
16 thought, language is yet another domain in which the brain acts as a prediction machine
17 (Clark, 2013; Van Berkum, 2010; see also Friston, 2005, 2010; Summerfield & De Lange,
18 2014), hard-wired to continuously match sensory inputs with top-down, grammatical or
19 probabilistic expectations based on context and memory.

20 What promoted linguistic prediction from outlandish and deeply contentious to
21 ubiquitous and somewhat anodyne? One of the key and most compelling pieces of empirical
22 evidence for linguistic prediction to date comes from a landmark Nature Neuroscience
23 publication by DeLong, Urbach and Kutas (2005), whose approach exploited a phonological
24 rule of English whereby the indefinite article is realized as a before consonant-initial words
25 and as an before vowel-initial words. In their experiment, participants read sentences of

26 varying degree of contextual constraint that led to expectations for a particular consonant- or
27 vowel-initial noun. This expectation was operationalized as a word's cloze probability
28 (cloze), calculated in a separate, non-speeded sentence completion task as the percentage of
29 continuations of a sentence fragment with that word (Taylor, 1953). For example, the
30 sentence fragment "The day was breezy so the boy went outside to fly..." is continued with
31 'a' by 86% of participants, and "The day was breezy so the boy went outside to fly a...", it is
32 continued with 'kite' by 89% of participants. In the main experiment, word-by-word sentence
33 presentation enabled DeLong and colleagues to examine electrical brain activity elicited by
34 articles that were concordant with the highly expected but yet unseen noun ('a', followed by
35 'kite'), or by articles that were incompatible with the highly expected noun and heralded a
36 less expected one ('an', followed by 'airplane'). The dependent measure was the amplitude of
37 the N400¹ event-related potential (ERP), a negative ERP deflection that peaks approximately
38 400 ms after word onset and is maximal at centroparietal electrodes (Kutas & Hillyard,
39 1980). The N400 is elicited by every word of an unfolding sentence and its amplitude is
40 smaller (less negative) with increasing ease of semantic processing (Kutas & Hillyard, 1984).
41 DeLong et al. found that the N400 amplitude for a given word decreased as a function of
42 increasing cloze probability, both for nouns and, critically, for articles. The systematic,
43 graded N400 modulation by article-cloze was taken as strong evidence that participants
44 activated the nouns in advance of their appearance, and that the disconfirmation of this

¹ In this article, we use "N400 amplitude" as a shorthand for "ERP amplitude in the time window associated with the N400"; this ERP amplitude is actually a sum of the N400 ERP component and other ERP components (reflecting other aspects of cognition) that overlap with it in time and space.

45 prediction by the less-expected articles resulted in processing difficulty (higher N400
46 amplitude at the article).

47 The results obtained with this elegant design warranted a much stronger conclusion
48 than related results available at the time. Previous studies that employed a visual-world
49 paradigm had revealed listeners' anticipatory eye-movements towards visual objects on the
50 basis of probabilistic or grammatical considerations (e.g., Altmann & Kamide, 1999).
51 However, predictions in such studies are scaffolded onto already-available visual context, and
52 therefore do not measure purely pre-activation, but perhaps re-activation of word information
53 previously activated by the visual object itself (Huettig, 2015). DeLong and colleagues
54 examined brain responses to information associated with concepts that were not pre-specified
55 and had to be retrieved from long-term memory 'on-the-fly'. Furthermore, DeLong and
56 colleagues were the first to muster evidence for highly specific pre-activation of a word's
57 phonological form, rather than merely its semantic (e.g., Federmeier & Kutas, 1999) or
58 morpho-syntactic features (e.g., Van Berkum, Brown, Zwitserlood, Kooijman & Hagoort,
59 2005; Wicha, Moreno & Kutas, 2004). Crucially, as their demonstration involved
60 semantically identical articles (function words) rather than nouns or adjectives (content
61 words) that are rich in meaning, the observed N400 modulation by article-cloze is unlikely to
62 reflect difficulty interpreting the articles themselves. Most notably, DeLong and colleagues
63 were the first to examine brain activity elicited by a range of more- or less-predictable
64 articles, not simply most- versus least-expected. Based on the observed correlation, they
65 argued that pre-activation is not all-or-none and limited to highly constraining contexts, but
66 occurs in a graded, probabilistic fashion, with the strength of a word pre-activation
67 proportional to its cloze probability. Moreover, they concluded that prediction is an integral
68 part of real-time language processing and, most likely, a mechanism for propelling the
69 comprehension system to keep up with the rapid pace of natural language.

70 DeLong et al.'s study has had an immense impact on psycholinguistics,
71 neurolinguistics and beyond. It is cited by authoritative reviews (e.g., Altmann & Mirkovic,
72 2009; Hagoort, 2017; Lau, Phillips & Poeppel, 2008; Pickering & Clark, 2014; Pickering &
73 Garrod, 2007) as delivering decisive evidence for probabilistic prediction of words all way up
74 to their phonological form. Moreover, as a demonstration of pre-activation of phonological
75 form (sound) during reading, it is often cited as evidence for 'prediction through production'
76 (e.g., Pickering & Garrod, 2013), the hypothesis that linguistic predictions are implicitly
77 generated by the language production system. To date, DeLong et al. has received a total of
78 735 citations (Google Scholar), averaging to more than 1 citation per week over the past
79 decade, with an increasing number of citations in each subsequent year. The results also
80 played an important role in settling an ongoing debate in the neuroscience of language. It
81 provided the clearest evidence that the N400 component, which for 25 years had been taken
82 to directly index the high-level compositional processes by which people integrate a word's
83 meaning with its context (Brown & Hagoort, 1993; Chwilla, Brown & Hagoort, 1995;
84 Connolly & Phillips, 1994; Friederici, Steinhauer & Frisch, 1999; Van Berkum, Hagoort &
85 Brown, 1999; Van Petten, Coulson, Rubin, Plante, & Parks, 1999), actually reflected non-
86 compositional processes by which word information is accessed as a function of context.

87 But how robust are gradient effects of form prediction? In over a decade that has
88 passed since the publication by DeLong and colleagues, there is still no published study that
89 directly replicates their graded pattern of results (for an overview, see Ito, Martin &
90 Nieuwland, 2017b). An alternative analysis of the same data by the authors did not yield a
91 statistically significant result (DeLong, 2009), but was not mentioned in the published report.
92 In at least three other unpublished data sets (DeLong, 2009; Miyamoto, 2016), DeLong and
93 colleagues did not find a significant correlation between article-N400 and cloze probability.
94 Martin, Thierry, Kuipers, Boutonnet, Foucart and Costa (2013) reported a successful

95 conceptual replication in native speakers of English but not in bilinguals. However, their
96 study did not test for a graded effect of cloze, and differed from the original in many crucial
97 aspects of the experimental design, data-preprocessing and statistical analysis, clouding both
98 a qualitative and quantitative comparison to the original results. Moreover, two attempts to
99 replicate the Martin et al. results in English monolinguals failed to yield a reliable effect of
100 cloze on article-ERPs (Ito, Martin & Nieuwland, 2017a,b).

101 As the tremendous scientific impact of the DeLong et al. findings is at odds with the
102 apparent lack of replication attempts, we report here a direct replication study. Inspired by
103 recent demonstrations for the need for large subject-samples in psychology and neuroscience
104 research (Button et al., 2013; Open Science Collaboration, 2015), our replication spanned 9
105 laboratories each with a sample size equal to or greater than that of the original. In addition to
106 duplicating the original analysis, our replication attempt also seeks to improve upon DeLong
107 et al.'s data analysis. DeLong et al.'s original analysis reduced an initial pool of 2560 data
108 points (32 subjects who each read 80 sentences) to 10 grand-average values, by averaging
109 N400 responses over trials within 10 cloze probability decile-bins (cloze 0-10, 11-20, et
110 cetera), per participant and then averaging over participants, even though these bins held
111 greatly different numbers of observations (for example, the 0-10 cloze bin contained 37.5%
112 of all data). These 10 values were correlated with the average cloze value per bin, yielding
113 numerically high correlation coefficients with large confidence intervals (for example, the Cz
114 electrode showed a statistically significant r-value of 0.68 with a 95% confidence interval
115 ranging from 0.09 to 0.92). However, this analysis potentially compromises power by
116 discretizing cloze probability into deciles and not distinguishing various sources of subject-,
117 item-, bin-, and trial-level variation. Furthermore, treating subjects as fixed rather than
118 random factor potentially inflates false positive rates, since the overall cloze effect is

119 confounded with by-subject variation in the effect (Barr, Levy, Scheepers & Tily, 2013;
120 Clark, 1973).

121 In our replication study, we followed two pre-registered analysis routes: a replication
122 analysis that duplicated the DeLong et al. analysis, and a single-trial analysis that modelled
123 variance at the level of item and subject. The effect of cloze on noun-elicited N400s (DeLong
124 et al., 2015; Kutas & Hillyard, 1984) is a necessary but not sufficient evidence for the claim
125 on pre-activation in language processing (as it is also compatible with the view that the
126 noun's cloze probability correlates with the ease of integration of that noun into the context).
127 It serves as a manipulation check to ensure that the experiment is able to successfully detect
128 graded variation in N400 amplitude, but does not provide strong evidence for the prediction
129 of phonological form. That evidence would come from the ERPs elicited by articles.
130 Observing a reliable effect of cloze on article-elicited N400s in the replication analysis and,
131 in particular, in the single-trial analysis, would constitute powerful evidence for the pre-
132 activation of phonological form during reading.

133

134 **RESULTS**

135 We first obtained offline cloze probabilities for all target articles and nouns from a
136 group of native English speakers. These values closely resembled those of the original study
137 (see Methods for details). In the subsequent ERP experiment, a different group of participants
138 (N=334) read the sentences word-by-word from a computer display at a rate of 2 words per
139 second while we recorded their electrical brain activity at the scalp. The replication analysis
140 and single-trial analysis described below were each pre-registered at <https://osf.io/eyzaq/>.

141 **Replication analysis**

142 We sorted the articles and nouns into 10 bins based on each word's cloze probability
143 (e.g., items with 0-10% cloze were put in one bin, 10-20% in another, etc.). For each

144 laboratory, we averaged ERPs per bin first within, then across, participants. No baseline
145 correction was used, following the procedure described in the Methods section in DeLong et
146 al (2005). We then correlated the averaged cloze values per bin with mean ERP amplitude in
147 the N400 time window (200-500 ms) elicited by the nouns (for the noun analysis) or articles
148 (for the article analysis) from the corresponding bin, yielding a Pearson correlation coefficient
149 (r-value) per EEG channel. This analysis yielded a very different pattern than DeLong et al.
150 observed (Fig. 1). In no laboratory did article-N400 amplitude at centro-parietal sites become
151 significantly smaller (less negative) as article-cloze probability increased (in fact, in most
152 laboratories the pattern went into the opposite direction). Only in one laboratory (Lab 2) did
153 the correlation coefficient have a p-value below .05 in the predicted direction (positive) at
154 any electrode (uncorrected for multiple comparisons), but this effect was observed at a few
155 left-frontal electrodes, not at the central-parietal electrodes where DeLong et al found their
156 N400 effects. Moreover, in two laboratories (Labs 3 and 5), a statistically significant effect
157 was observed in the opposite direction, larger (more negative) article-N400 amplitude for
158 articles with increasing cloze probability. For the nouns, the pattern was more similar to the
159 DeLong et al. results. In six laboratories (Lab 2, 3, 4,6, 7, and 9), noun-N400 amplitude for
160 nouns at central-parietal or parietal-occipital electrodes became smaller with increasing noun-
161 cloze, and in two other laboratories (Lab 5 and 8) the effects clearly went in the expected
162 direction without reaching statistical significance.

163 DeLong et al. recently mentioned using a 500 ms baseline correction procedure that
164 was not mentioned in the published study (personal communication by DeLong, March
165 2017). In an exploratory analysis, we therefore recomputed the correlations based on data
166 pooled from all laboratories using this baseline correction procedure (Fig 2.). This analysis
167 also showed a lack of statistically significant positive correlations for the articles, but
168 statistically significant positive correlations for the nouns. In exploratory Bayesian analyses

169 reported below, we perform an analysis to establish whether these results are consistent with
170 the size and direction of the effects reported by DeLong et al., regardless of statistical
171 significance.

172

173 **Single-trial analysis**

174 We first performed baseline correction by subtracting the average amplitude in the 100
175 ms time window before word onset. Baseline-corrected ERPs for relatively expected and
176 unexpected words and difference waveforms are shown in Fig. 3. Then, for the data pooled
177 across all laboratories, we used linear mixed effects models to regress the N400 amplitude (in
178 a spatiotemporal region of interest selected a priori based on the DeLong et al. results) on
179 cloze probability. For the articles, the effect of cloze was not statistically significant at the
180 $\alpha=.05$ level, $\beta = .29$, CI [-.08, .67], $\chi^2(1) = 2.31$, $p = .13$ (see Fig. 4, left panel)², with β
181 referring to the N400 difference in microvolts associated with stepping from 0% to 100%
182 cloze. The effect of cloze on N400 amplitude at the article did not significantly differ
183 between laboratories, $\chi^2(8) = 7.90$, $p = .44$. For the nouns, however, higher cloze values were
184 strongly associated with smaller N400s, $\beta = 2.22$, CI [1.76, 2.69], $\chi^2(1) = 56.50$, $p < .001$ (see
185 Fig. 4, right panel). This pattern did not significantly differ between laboratories, $\chi^2(8) =$
186 11.59, $p = .17$. The effect of cloze on noun-N400s was statistically different from its effect on
187 article-N400s, $\chi^2(1) = 31.38$, $p < .001$.

188 Exploratory (i.e., not pre-registered) single-trial analyses

189 The effect of article-cloze did not significantly vary as a function of subject
190 comprehension question accuracy, $\chi^2(1) = 0.45$, $p = .50$. In addition, the effect of article-cloze

² Unless otherwise indicated, p-values are two-tailed, and CIs are two-tailed 95% confidence intervals.

191 was also not statistically significant when subject comprehension accuracy was included in
192 the analysis (100 ms baseline: $\beta = .24$, CI [-.17, .64], $\chi^2(1) = 1.27$, $p = .26$).

193 In our dataset, an analysis in the 500 to 100 ms time window before article-onset revealed
194 a non-significant effect of cloze that resembled the pattern observed after article-onset, $\beta =$
195 $.16$, CI [-.07, .39], $\chi^2(1) = 1.82$, $p = .18$. This suggested that a 500 ms baseline correction
196 procedure, which was used but not reported in DeLong et al. (2005), would better correct for
197 pre-article voltage-levels. We repeated our analysis with the 500 ms baseline correction
198 procedure, the initially observed effect of article-cloze was numerically smaller and less
199 significant than it was in the pre-registered analysis ($\beta = .14$, CI [-.25, .53], $\chi^2(1) = 0.46$, $p =$
200 $.50$).

201

202 **Exploratory Bayesian analyses**

203 For the articles, our pre-registered replication analyses yielded non-significant p -
204 values, indicating failure to reject the null-hypothesis that cloze has no effect on N400
205 activity. To better adjudicate between the null-hypothesis (H_0) and an alternative hypothesis
206 (H_r), we performed exploratory replication Bayes factor analysis for correlations
207 (Wagenmakers, Verhagen & Ly 2016). The obtained replication Bayes factor quantifies the
208 evidence that there is an effect in the size and direction reported by DeLong et al. (see Fig. 5).
209 For the articles, this yielded strong to extremely strong evidence for the null hypothesis that
210 the effect of cloze is zero, with BF_{0r} values up to 154 (at the Cz electrode depicted by
211 DeLong et al., $BF_{0r} = 77$), and strongest evidence at the posterior channels. For the nouns, we
212 obtained extremely strong evidence for the alternative hypothesis that the effect is nonzero,
213 particularly at posterior channels, with BF_{10} values up to 9,163,515 (at Cz, $BF_{r0} = 10,725$).
214 The pattern of results was similar when the 500 ms pre-stimulus baseline was applied.

215 Next, we computed Bayesian mixed-effect model estimates (β) and 95% credible
216 intervals (CrI) for our single-trial analyses, using priors based on the results from DeLong et
217 al. In both of our article-analyses credible intervals included zero (100 ms baseline: $\beta = .31$,
218 CrI [-.06 .69]; 500 ms baseline: $\beta = .17$, CrI [-.22 .55]). For the nouns, zero was not within
219 the credible interval: $\beta = 2.24$, CrI [1.77 2.70]. These Bayesian analyses further demonstrate
220 our failure to replicate the DeLong et al. article-effect alongside a successful replication of
221 the noun-effect. The analyses suggest that the data (combined with prior assumptions about
222 the effect) are not very consistent with the hypothesis that the article-effect is zero (further
223 information and posterior summaries are available in Supplementary Figure 2), but also are
224 extremely inconsistent with the hypothesis that the article-effect is as big as that observed by
225 DeLong and colleagues (2005). The data are most consistent with an effect that is more likely
226 to be positive than zero or negative, but is very small (so small that it was not detected at
227 traditional significance levels in this large-scale experiment with substantially higher power
228 than previous experiments).

229

230 **Control experiment**

231 Lack of a statistically significant, article-elicited prediction effect could reflect a
232 general insensitivity of our participants to the phonologically conditioned variation of the
233 English indefinite article, i.e., a/an alternation. We ruled out this alternative explanation in an
234 additional experiment that followed the replication experiment as part of the same
235 experimental session. Participants read 80 short sentences containing the same nouns as the
236 replication experiment, preceded by a phonologically licit or illicit article (e.g., “David found
237 a/an apple...”), presented in the same manner as before. In each laboratory, nouns following
238 illicit articles elicited a late positive-going waveform compared to nouns following licit
239 articles (see Fig. 6), starting at about 500 ms after word onset and strongest at parietal

240 electrodes. This standard P600 effect (Osterhout & Holcomb, 1992) was confirmed in a
241 single-trial analysis, $\chi^2(1) = 83.09$, $p < .001$, and did not significantly differ between labs,
242 $\chi^2(8) = 8.98$, $p = .35$.

243 **DISCUSSION**

244 In a landmark study, DeLong, Urbach and Kutas observed a statistically significant,
245 graded modulation of article- and noun-elicited electrical brain potentials (N400) by the pre-
246 determined probability that people continue a sentence fragment with that word (cloze). They
247 concluded that people routinely and probabilistically pre-activate upcoming words to a high
248 level of detail, including whether a word starts with a consonant or vowel. Our direct
249 replication study spanning 9 laboratories successfully replicated a statistically significant
250 effect of cloze on noun-elicited N400 activity but, critically, failed to replicate such an effect
251 of cloze on article-elicited N400 activity. This pattern of success and failure was observed in
252 a pre-registered replication analysis that duplicated the original study's analysis, and a pre-
253 registered single-trial analysis that modelled variance at the level of item and subject.
254 Exploratory Replication Bayes Factor analyses confirmed that we successfully replicated the
255 direction and size of the correlations reported by DeLong et al. for the nouns, but not for the
256 articles. Exploratory Bayesian mixed-effects model analyses suggested that, while there is
257 some evidence that the true population-level effect may be in the direction reported by
258 DeLong and colleagues, the effect is likely far smaller than what they reported. In fact, the
259 effect is likely too small to be meaningfully observed without very large sample sizes,
260 hence of uncertain theoretical interest. Finally, a control experiment confirmed that our
261 participants did respect the phonological alternation a/an of the article with nouns used in the
262 replication experiment.

263 Our findings carry important theoretical implications by challenging a crucial cornerstone
264 of the ‘strong prediction view’ held by current theories of language comprehension (e.g.,
265 Altmann & Mirkovic, 2009; Pickering & Garrod, 2013). The strong prediction view entails
266 two key claims. The first is that people pre-activate words at all levels of representation in a
267 routine and implicit (i.e., non-strategic) fashion. Pre-activation is not limited to a word’s
268 meaning, but includes its grammatical features and even its orthographic and/or phonological
269 form. This would put language on a par with other cognitive systems such as visual
270 perception that attempt to predict the inputs to lower-level ones (Friston, 2005, 2010;
271 Summerfield & De Lange, 2014). The second claim is that pre-activation occurs at all levels
272 of contextual support and gradually increases in strength with the level of contextual support.
273 When contextual support for a specific word is high, like at a 100% cloze value, the word’s
274 form and meaning is strongly pre-activated. When contextual support for a word is low, like
275 when it is one amongst 20 words each with a 5% cloze value, pre-activation is distributed
276 across multiple potential continuations. However, even then, a word’s form and meaning are
277 pre-activated, just weakly so. The strength of pre-activation is probabilistic, that is, linked to
278 estimated probability of occurrence.

279 DeLong and colleagues, and subsequently other scientists (e.g., Dell & Chang, 2014;
280 Pickering & Clark, 2013), took their results as the evidence to support both these claims.
281 DeLong et al (2005) was – and still is - the only study to date that measured pre-activation at
282 the prenominal articles *a* and *an* that do not differ in their semantic or grammatical content,
283 and that observed a graded relationship between cloze and N400 activity across a range of
284 low- and high-cloze words, rather than merely a difference between low- and high-cloze
285 words. Given that the use of these articles depends on whether the next word starts with a
286 vowel or consonant, their results were considered as powerful evidence that participants
287 probabilistically pre-activated the initial sound of upcoming nouns.

288 However, we show that there is no statistically significant effect of cloze on article-
289 elicited N400 activity, using a sample size more than ten times that of the original, and a
290 statistical analysis that better accounts for sources of non-independence than the original
291 averaging-based correlation approach. If an effect of cloze on article-N400s exists at all, its
292 true effect size is so small that it cannot be reliably detected even in an expansive multi-
293 laboratory approach, let alone in the typical sample size in psycholinguistic and
294 neurolinguistic experiments (roughly, $N= 30$). This means that even if article-cloze is
295 associated with a graded modulation of N400 amplitudes, this effect seems to be so small that
296 it cannot be reliably measured with small samples, and thus the previous studies may not
297 have contributed much reliable information to our understanding of this effect. Moreover, it
298 is also possible that the effect is sensitive to specifics of the experimental procedure and
299 context that it lacks generalizability. Current theoretical positions thus either require new
300 strong evidence for phonological pre-activation or require revision. In particular, the strong
301 prediction view that claims that pre-activation routinely occurs across all – including
302 phonological – levels (Pickering & Garrod, 2013), can no longer be viewed as having strong
303 empirical support. Our work impels the field think differently about what constitutes strong
304 evidence within a theory, but also highlights the need for a theory of linguistic prediction to
305 formulate quantitative predictions about the effect-size of to-be-observed effects.

306 By contrast, we observed a strong and statistically significant effect of cloze on noun-
307 elicited activity in the majority of our analyses. Although three of the nine laboratories did
308 not show statistically significant correlations between noun-cloze and N400s, data pooled all
309 laboratories showed a strong and statistically significant noun-cloze effect, our Replication
310 Bayes Factor analysis overwhelmingly replicated the direction and size of the noun-cloze
311 effect of DeLong et al., and our more powerful single-trial analysis revealed a significant
312 noun-cloze effect in each of the laboratories. These results are therefore consistent with the

313 handful of studies that reported a graded relationship between noun-cloze and noun-N400s
314 (DeLong et al., 2005; Kutas & Hillyard, 1984; Wlotko & Federmeier, 2012).

315 Where does this pattern of failure and success leave the strong prediction view?
316 Following the experimental logic of DeLong et al, we do not have sufficient evidence to
317 conclude that people routinely pre-activate the initial phoneme of an upcoming noun, or
318 perhaps any other word form information. Without pre-activation of the initial phoneme, the
319 specific instantiation of the article does not cause people to revise their prediction about the
320 meaning of the upcoming noun, thus lacking any impact on processing. Crucially, this
321 conclusion is incompatible with the strong prediction view, because it suggests that pre-
322 activation does not occur to the level of detail that is often assumed. Our results are also
323 incompatible with an alternative interpretation of the DeLong et al. findings that people
324 predict the article itself together with the noun (Ito, Corley, Pickering, Martin & Nieuwland,
325 2016; Van Petten & Luka, 2012), and they pose a serious challenge to the theory that
326 comprehenders predict upcoming words, including their initial phonemes, through implicit
327 production (Pickering & Garrod, 2013). Crucially, the idea that prediction is probabilistic,
328 rather than all-or-none, is now questionable, given that there is no other published report of a
329 pre-activation gradient. Although other studies have claimed prediction of form (Ito et al.,
330 2016) or a prediction gradient (Smith & Levy, 2013), no study has indisputably demonstrated
331 graded pre-activation, i.e., graded effects occurring before the noun. Effects that are observed
332 upon, rather than before the noun, do not purely index pre-activation but index a mixture of
333 memory retrieval and semantic integration processes instigated by the noun itself (Baggio &
334 Hagoort, 2011; Lau, Namyst, Fogel & Delgado, 2016; Otten & Van Berkum, 2008;
335 Steinhauer, Royle, Drury & Fromont, 2017). Therefore, there is currently no clear evidence to
336 support routine probabilistic pre-activation of a noun's phonological form during sentence
337 comprehension.

338 Our results, however, do not necessarily exclude phonological form pre-activation, and
339 we temper our conclusion with a caveat stemming from the a/an manipulation. For this
340 manipulation to ‘work’, people must specifically predict the initial phoneme of the next word,
341 and revise this prediction when faced with an unexpected article. However, because articles
342 are only diagnostic about the next word within the noun phrase, rather than about the head
343 noun itself, an unexpected article does not refute the upcoming noun, it merely signals that
344 another word would come first (e.g., ‘an old kite’). This opens up explanations for why the
345 a/an manipulation ‘fails’. In addition, comprehenders may not predict the noun to follow
346 immediately, but at a later point; the unexpected article then does not evoke a change in
347 prediction. Predictions about a specific position may be disconfirmed too often in natural
348 language to be viable. This idea is supported by corpus data (Corpus of Contemporary
349 American English and British National Corpus.), showing a mere 33% probability that a/an is
350 directly followed by a noun. Alternatively, people predict the noun to come next, but only
351 revise their prediction about its linear position while retaining the prediction about its
352 meaning. So perhaps a revision of the predicted meaning, not the position, is required to
353 trigger differential ERPs. In both of these hypothetical scenarios, people do not revise their
354 prediction about the upcoming noun’s meaning unless they must.

355 Our results can be straightforwardly reconciled with effects reported for other pre-
356 nominal manipulations, such as those of Dutch or Spanish article-gender (e.g., Van Berkum
357 et al., 2008; Otten, Nieuwland & Van Berkum, 2008; Otten & Van Berkum, 2009; Wicha et
358 al., 2004). Unlike a/an articles, gender-marked articles can immediately disconfirm the noun,
359 because article- and noun-gender agrees regardless of intervening words (e.g., the Spanish
360 article ‘el’ heralds a masculine noun). Revising the prediction about the noun presumably
361 results in a semantic processing cost, thereby modulating N400 activity. Although gender-
362 marked articles do not consistently incur the exact same type of effect and have only been

363 observed at very high cloze values, previous studies suggest that a noun's grammatical
364 gender can be pre-activated along with its meaning. Compared to this gender-manipulation,
365 DeLong et al's study based on the English a/an manipulation claimed a stronger version of
366 the prediction view, namely that people predict which word comes next up to its phonological
367 form and, make backwards prediction as to the phonological form of the preceding linguistic
368 material even on the basis of probabilistic, graded information.

369 What do our results say about prediction during natural language processing? Like the
370 conclusions by DeLong et al., ours are limited by the generalization from language
371 comprehension in a laboratory setting. On one hand, a rich conversational or story context
372 may enhance predictions of upcoming words, and listeners may be more likely to pre-activate
373 the phonological form of upcoming words than readers. On the other hand, our laboratory
374 setting offered particularly good conditions for prediction of the next word's initial sound to
375 occur. Each article was always immediately followed by a noun, unlike in natural language.
376 Moreover, compared to natural reading rates our word presentation rate was slow, which may
377 facilitate predictive processing (Ito et al., 2016; Wlotko & Federmeier, 2015). In natural
378 reading, articles are hardly fixated and often skipped (e.g., O'Regan 1979). In short,
379 arguments can be made both for and against phonological form prediction in natural language
380 settings, and novel avenues of experimentation are needed to settle this issue.

381 DeLong and colleagues recently stated an omission in the description of their data
382 analysis, i.e., a baseline procedure was applied to the data but inadvertently omitted from the
383 description (DeLong et al., 2005). We have shown that our conclusions hold regardless of the
384 baseline procedure. In a recent commentary, DeLong, Urbach, and Kutas (2017) also
385 described filler-sentences in their experiment, which were omitted from their original report,
386 and were neither provided nor mentioned to us by the authors upon our request for the
387 stimuli. DeLong et al. used the existence of these filler-sentences to dismiss an alternative

388 explanation of their original findings, namely that an unusual experimental context wherein
389 every sentence contains an article-noun combination leads participants to strategically predict
390 upcoming nouns. Following this logic, we failed to replicate their article-effects despite an
391 experimental context that could inadvertently encourage strategic prediction (for
392 demonstrations of experimental context boosting predictive processing, see Brothers, Swaab
393 & Traxler, 2017; Lau, Holcomb & Kuperberg, 2013). Therefore, the presence of fillers in
394 their experiment versus absence in ours cannot straightforwardly explain the different results,
395 and may even strengthen our conclusions.

396 To conclude, we failed to replicate the main result of DeLong et al., a landmark study
397 published more than ten years ago that has not been directly replicated since. Our results
398 suggest that, if there is an effect of article-cloze probability on the amplitude of the N400, it
399 is too small and/or too sensitive to unknown experimental design factors to have been
400 meaningfully measured in previous small-sample-size experiments. We conclude that such an
401 effect does not constitute strong evidence for current theoretical positions on the importance
402 of prediction (e.g., Pickering & Garrod, 2013). Our findings thus challenge one of the pillars
403 of the ‘strong prediction view’ in which people routinely and probabilistically pre-activate
404 information at all levels of linguistic representation, including phonological form information
405 such as the initial phoneme of an upcoming noun. Consequently, there is currently no
406 convincing evidence that people pre-activate the phonological form of an upcoming noun
407 during sentence comprehension, and we take our findings to suggest a more limited role for
408 prediction during language comprehension. In addition, our findings further highlight the
409 importance of direct replication, large sample size studies, transparent reporting and of pre-
410 registration to advance reproducibility and replicability in the neurosciences.

411 MATERIALS AND METHODS

412 **Experimental design and materials.** Nieuwland twice requested all original
413 materials from DeLong et al., including the questions and norms, with the stated purpose of
414 direct replication (personal communication, November 4 and 19, 2015), upon which DeLong
415 et al. made available the 80 sentences described in the original study. These sentences were
416 then adapted from American to British spelling and underwent a few minor changes to ensure
417 their suitability for British participants. The complete set of materials and the list of changes
418 to the original materials are available online (Supplementary Table 1 and 2). The materials
419 were 80 sentence contexts with two possible continuations each: a more or less expected
420 indefinite article + noun combination. The noun was followed by at least one subsequent
421 word. All article + noun continuations were grammatically correct. Each article + noun
422 combination served once as the more expected continuation and the other time as the less
423 expected continuation, in different contexts. We divided the 160 items in two lists of 80
424 sentences such that each list contained each noun only once. Each participant was presented
425 with only one list (thus, each context was seen only once). One in four sentences was
426 followed by a yes/no comprehension question, which yielded a mean response accuracy of
427 95% (after taking into account ambiguity in three of the questions, see Supplemental Table 2
428 and 3). While this percentage is very similar to that reported by DeLong et al., we note that
429 this cannot be directly compared to the accuracy reported in DeLong et al., because we had to
430 create new comprehension questions in the absence of the original ones. Regardless, because
431 DeLong et al. suggested that our results were due to poor language comprehension (DeLong,
432 Urbach & Kutas, 2017), we describe an exploratory analysis in which we attempt to account
433 for variation in response accuracy in the statistical model.

434 We obtained article cloze and noun cloze ratings from a separate group of native
435 speakers of English who were students at the University of Edinburgh and did not participate

436 in the ERP experiment. They were instructed to complete the sentence fragment with the best
437 continuation that comes to mind (Taylor, 1953). We obtained article cloze ratings from 44
438 participants for 80 sentence contexts truncated before the critical article. Noun cloze ratings
439 were obtained by first truncating the sentences after the critical articles, and presenting two
440 different, counterbalanced lists of 80 sentences to 30 participants each, such that a given
441 participant only saw each sentence context with the expected or the unexpected article. The
442 obtained values closely resemble those of the original study, with the same range (0-100% for
443 articles and nouns), slightly lower median values (for articles and nouns, 29% and 40%,
444 compared to 31% and 46% in the original study), but slightly higher mean values (for articles
445 and nouns, 41% and 46%, compared to 36% and 44%). Because the sentence materials we
446 used describe common situations that can be understood by any English speaker, and because
447 students at the University of Edinburgh come from across the whole of the UK, we had no a
448 priori expectation that cloze ratings would differ substantially across laboratories, and thus
449 we did not obtain cloze norms from other sites. Consistent with this assumption, nothing in
450 our results suggests stronger cloze effects in University of Edinburgh students compared to
451 other students, suggesting that our cloze norms are sufficiently representative for the other
452 universities.

453 **Participants.** Participants were students from the University of Birmingham, Bristol,
454 Edinburgh, Glasgow, Kent, Oxford, Stirling, York, or volunteers from the participant pool of
455 University College London or Oxford University, who received cash or course credit for
456 taking part in the ERP experiment. Participant information and EEG recording information
457 per laboratory is available online (Supplementary Table 3). We pre-registered a target sample
458 size of 40 participants per laboratory, which was thought to give at least 32 participants (the
459 sample size of DeLong et al.) per laboratory after accounting for data loss, as was later
460 confirmed. Due to logistic constraints, not all laboratories reached an N of 40. Because in two

461 labs corruption of data was incorrectly assumed before computing trial loss, these
462 laboratories tested slightly more than 40 participants. All participants (N = 356; 222 women)
463 were right-handed, native English speakers with normal or corrected-to-normal vision,
464 between 18–35 years (mean, 19.8 years), free from any known language or learning disorder.
465 Eighty-nine participants reported a left-handed parent or sibling.

466 **Procedure.** After giving written informed consent, participants were tested in a single
467 session. Sentences were presented visually in the center of a computer display, one word at a
468 time (200 ms duration, followed by a blank screen of 300 ms duration³). Participants were
469 instructed to read sentences for comprehension and answer yes/no comprehension questions
470 by pressing hand-held buttons. The electroencephalogram (EEG) was recorded from at least
471 32 electrodes.

472 The replication experiment was followed by a control experiment, which served to
473 detect sensitivity to the correct use of the a/an rule in our participants. Participants read 80
474 relatively short sentences (average length 8 words, range 5-11) that contained the same
475 critical words as the replication experiment, preceded by a correct or incorrect article. As in

³ Due to a programming error, in four labs (1, 3, 5 and 8, which used E-prime scripts) the critical articles and nouns, but not other words, were followed by a 380 ms blank instead of the intended 300 ms. However, this delay is unlikely to have affected the results because if it was noticed at all, which is unlikely, it appeared after the N400 window associated with the article. Moreover, if anything, longer duration facilitates language comprehension and predictive processing (Camblin, LeDoux, Boudewyn, Gordon & Swaab, 2007; Wlotko & Federmeier, 2015; Ito et al., 2016), making it more, not less likely to find an effect of cloze on the article-ERPs. Of note, the qualitative pattern of the results from the pre-registered single-trial analysis did not change when we removed these labs from the analysis.

476 the replication experiment, each critical word was presented only once, and was followed by
477 at least one more word. All words were presented at the same rate as the replication
478 experiment. There were no comprehension questions in this experiment. After the control
479 experiment, participants performed a Verbal Fluency Test and a Reading Span test; the
480 results from these tests are not discussed here. All stimulus presentation scripts are publicly
481 available in two different software packages (E-Prime and Presentation) on
482 <https://osf.io/eyzaq>.

483 **Data processing.** Data processing was performed in BrainVision Analyzer 2.1 (Brain
484 Products, Germany). We performed one pre-registered replication analysis that followed the
485 DeLong et al. analysis as closely as possible and one pre-registered single-trial analysis
486 (Open Science Framework, <https://osf.io/eyzaq>). All non-pre-registered analyses are
487 considered as exploratory. First, we interpolated bad channels from surrounding channels,
488 and downsampled to a common set of 22 EEG channels per laboratory which were similar in
489 scalp location to those used by DeLong et al. One laboratory did not have 12 of the selected
490 22 channels in its EEG channel montage, and we matched the full 22-channel layout used for
491 other laboratories by creating 12 virtual channels from neighbouring channels using
492 topographic interpolation by spherical splines. We then applied a 0.01-100 Hz digital band-
493 pass filter (including 50 Hz Notch filter), re-referenced all channels to the average of the left
494 and right mastoid channels (in a few participants with a noisy mastoid channel, only one
495 mastoid channel was used), and segmented the continuous data into epochs from 500 ms
496 before to 1000 ms after word onset. We then performed visual inspection of all data segments
497 and rejected data with amplifier blocking, movement artifacts, or excessive muscle activity.
498 Subsequently, we performed independent component analysis (Jung et al., 2000) on a 1-Hz
499 high-pass filtered version of the data, and applied the obtained weightings to the original data
500 to correct for blinks, eye movements or steady muscle artefacts. After this, we automatically

501 rejected segments containing a voltage difference of over 120 μV in a time window of 150
502 ms or containing a voltage step of over 50 $\mu\text{V}/\text{ms}$. Participants with fewer than 60/80 article
503 trials or 60/80 noun trials were removed from the analysis, leaving a total of 334 participants
504 (range across laboratories 32-42, and therefore each lab had a sample size at least as large as
505 DeLong et al.). On average, participants had 77 article trials and 77 noun trials.

506 **Pre-registered replication analysis.** We applied a 4th-order Butterworth band-pass
507 filter at 0.2-15 Hz to the segmented data, averaged trials per participant within 10% cloze
508 bins (0-10, 11-20, etc. until 91-100), and then averaged the participant-wise averages
509 separately for each laboratory. Because the bins did not contain equal numbers of trials (the
510 intermediate bins contained fewest trials), like in DeLong et al., not all participants
511 contributed a value for each bin to the grand average per laboratory. For nouns and articles
512 separately, and for each EEG channel, we computed the correlation between ERP amplitude
513 in the 200-500 ms time window per bin with the average cloze probability per bin.

514 **Pre-registered single-trial analysis.** In this analysis, we did not apply the 0.2-15 Hz
515 band-pass filter, which carries the risk of inducing data distortions (Luck, 2014; Tanner,
516 Morgan-Short & Luck, 2015). However, we deemed it necessary to perform a baseline
517 correction of the data. This procedure corrects for spurious voltage differences before word
518 onset, generating confidence that observed effects are elicited by the word rather than
519 differences in brain activity that already existed before the word and is a standard procedure
520 in ERP research (Luck, 2014). DeLong et al (2005) did not report a baseline correction, nor
521 did any of the related work from DeLong and colleagues that was reported in DeLong (2009).
522 Yet baseline correction has been used in many other publications from the Kutas Cognitive
523 Electrophysiology Lab. We chose a 100 ms pre-stimulus baseline as the most frequently used
524 one both in other studies from Kutas lab and in similar studies from other labs. For each trial,

525 we performed baseline correction by subtracting the mean voltage of the -100 to 0 ms time
526 window from each data point in the epoch.

527 Instead of averaging N400 data across trials and participants for subsequent statistical
528 analysis, we performed linear mixed effects model analysis (Baayen, Davidson & Bates,
529 2008) of the single-trial N400 data, using the “lme4” package (Bates, Maechler, Bolker &
530 Walker, 2014) in the R software (R CoreTeam, 2014). This approach simultaneously models
531 variance associated with each subject and with each item. Using a spatiotemporal region-of-
532 interest approach based on the DeLong et al. results, our dependent measure (N400
533 amplitude) was the average voltage across 6 centro-parietal channels (Cz/C3/C4/Pz/P3/P4) in
534 the 200-500 ms window for each trial. Analysis scripts and data to run these scripts are
535 publicly available on <https://osf.io/eyzaq>.

536 For articles and nouns separately, we used a maximal random effects structure as justified
537 by the design (Barr et al., 2013), which did not include random effects for ‘laboratory’ as
538 there were only 9 laboratories. Z-scored cloze was entered in the model as a continuous
539 variable, and laboratory was entered as a deviation-coded nuisance predictor. We tested the
540 effects of ‘laboratory’ and ‘cloze’ through model comparison with a χ^2 log-likelihood test.
541 We tested whether the inclusion of a given fixed effect led to a significantly better model fit.
542 The first model comparison examined laboratory effects, namely whether the cloze effect
543 varied across laboratories (cloze-by-laboratory interaction) or whether the N400 magnitudes
544 varied over laboratory (laboratory main effect). If laboratory effects were nonsignificant, we
545 dropped them from the analysis because they were not of theoretical interest. For the articles
546 and nouns separately, we compared the subsequent models below. Each model included the
547 random effects associated with the fixed effect ‘cloze’ (see Barr et al., 2014). All output β
548 estimates and 95% confidence intervals (CI) were transformed from z-scores back to raw

549 scores, and then back to the 0-100% cloze range, so that the voltage estimates represent the
550 change in voltage associated with a change in cloze probability from 0 to 100.

551 Model 1: $N400 \sim \text{cloze} * \text{laboratory} + (\text{cloze} | \text{subject}) + (\text{cloze} | \text{item})$

552 Model 2: $N400 \sim \text{cloze} + \text{laboratory} + (\text{cloze} | \text{subject}) + (\text{cloze} | \text{item})$

553 Model 3: $N400 \sim \text{cloze} + (\text{cloze} | \text{subject}) + (\text{cloze} | \text{item})$

554 Model 4: $N400 \sim (\text{cloze} | \text{subject}) + (\text{cloze} | \text{item})$

555 In an analysis that included the data from both articles and nouns, we also tested the
556 differential effect of cloze on article ERPs and on noun ERPs by comparing models with and
557 without an interaction between cloze and the deviation-coded factor ‘wordtype’
558 (article/noun). Random correlations were removed for the models to converge.

559 Model 1: $N400 \sim \text{cloze} * \text{wordtype} + (\text{cloze} * \text{wordtype} || \text{subject}) + (\text{cloze} * \text{wordtype} ||$
560 $\text{item})$

561 Model 2: $N400 \sim \text{cloze} + \text{wordtype} + (\text{cloze} * \text{wordtype} || \text{subject}) + (\text{cloze} * \text{wordtype} ||$
562 $\text{item})$

563 **Exploratory correlation analysis.** Of note, DeLong et al. have recently described
564 using a 500 ms baseline correction procedure that they failed to mention in DeLong et al.
565 (2005). Using this baseline correction procedure, we recomputed the correlations that we
566 obtained in our Replication analysis. To compare our results most directly with those reported
567 in Figure 1C of DeLong et al. (2005), we pooled data from all the laboratories so that we
568 would have a single r-value for each EEG-channel.

569 **Exploratory single-trial analyses.** We performed an exploratory analysis in the 500
570 to 100 ms time window before the article, using the originally (-100 to 0 ms) baselined data,
571 using Model 3 and 4 from the article analysis. This window covers the first 400 ms of the
572 word that preceded the article. Analysis in this window yielded a similar pattern as in the pre-
573 registered analysis, which indicates that a baseline correction procedure covering the entire

574 500 ms pre-stimulus window would account better for pre-article voltage levels. We
575 performed this additional analysis, the results of which did not change our conclusions and
576 are shown in Supplementary Figure 1.

577 We also performed an exploratory analysis in which we control for a potential
578 influence of response accuracy, taken as a proxy for the subject's attention to the task, on
579 predictive processing of linguistic input. We entered the (z-transformed) average response
580 accuracy of each subject in our model, and compared the models below. Comparison of
581 Model 1 and 2 tested whether the effect of cloze on the article-N400s depended subject
582 accuracy. Comparison of Model 2 and 3 tested whether there was a significant effect of cloze
583 on article-N400s when subject accuracy was included in the model.

584 Model 1: $N400 \sim \text{accuracy} * \text{cloze} + (\text{cloze} | \text{subject}) + (\text{cloze} | \text{item})$

585 Model 2: $N400 \sim \text{accuracy} + \text{cloze} + (\text{cloze} | \text{subject}) + (\text{cloze} | \text{item})$

586 Model 3: $N400 \sim \text{accuracy} + (\text{cloze} | \text{subject}) + (\text{cloze} | \text{item})$

587 **Exploratory Bayesian analyses.** Supplementing the Replication analysis, we
588 performed a Replication Bayes factor analysis for correlations (Wagenmakers et al., 2016)
589 using as prior the size and direction of the effect reported in the original study. We performed
590 this test for each electrode separately, after collapsing the data points from the different
591 laboratories. Because we had no articles in the 40-50 % cloze bin, there was a total of 9 and
592 10 data points per laboratory for the articles and nouns, respectively. Our analysis used priors
593 estimated from the DeLong et al results matched as closely as possible to our electrode
594 locations. A Bayes factor between 3 and 10 is considered moderate evidence, between 10-30
595 is considered strong evidence, 30-100 is very strong evidence, and values over 100 are
596 considered extremely strong evidence (Jeffreys, 1961). In addition to using a 100 ms pre-
597 stimulus baseline, we also computed the replication Bayes factors using the 500 ms pre-
598 stimulus time window for baseline correction. Results are shown in Figure 5.

599 Supplementing the pre-registered single-trial analyses, we performed an exploratory
600 Bayesian mixed-effects model analysis using the brms package for R (Buerkner, 2016),
601 which fits Bayesian multilevel models using the Stan programming language (Stan
602 Development Team, 2016). Nieuwland requested to use the results of a mixed-effects model
603 reanalysis of the DeLong et al. data as an appropriate prior (personal communication from
604 Nieuwland, November 14 and 22 2017); this request was declined by DeLong and
605 colleagues. We were therefore limited to using a prior centered on a point estimate based on
606 the DeLong et al. correlation analysis, namely our estimate of the observed effect size at Cz
607 for a difference between 0% cloze and 100% cloze (1.25 μ V and 3.75 μ V for articles and
608 nouns, respectively, based on visual inspection of the graphs) and a prior centered on zero for
609 the intercept. Both priors had a normal distribution and a standard deviation of 0.5 (given the
610 a priori expectation that average ERP voltages in this window generally fluctuate on the order
611 of a few microvolts; note that these units are expressed in terms of the z-scored cloze values,
612 rather than the original cloze values, such that μ for the cloze prior was 0.45, which
613 corresponds to a raw cloze effect of 1.25). We computed estimates and 95% credible intervals
614 for each of the mixed-effects models we tested, and transformed these back into raw cloze
615 units. The credible interval is the range of values such that one can be 95% certain that it
616 contains the true effect, given the data, priors and the model. The results from these analyses
617 are shown in Supplementary Figure 2; the analyses suggest that, while there may be a small
618 positive association between article cloze and ERP amplitude elicited by the articles, the
619 effect is substantially smaller than that estimated by DeLong and colleagues (2005) and likely
620 is too small to be meaningfully observed without very large sample sizes, hence of uncertain
621 theoretical interest.

622 **Control experiment.** Analysis of the control experiment involved a comparison
623 between a model with the categorical factor ‘grammaticality’ (grammatical/ungrammatical)

624 and a model without. Our dependent measure (P600 amplitude; Osterhout & Holcomb, 1992)

625 was the average voltage across 6 centro-parietal channels (Cz/C3/C4/Pz/P3/P4) in the 500-

626 800 ms window for each trial. Results are shown in Figure 6.

627 Model 1: $P600 \sim \text{grammaticality} + (\text{grammaticality} \mid \text{subject}) + (\text{grammaticality} \mid \text{item})$

628 Model 2: $P600 \sim (\text{grammaticality} \mid \text{subject}) + (\text{grammaticality} \mid \text{item})$

Acknowledgements

This work was partly funded by ERC Starting grant 636458 to H.J.F.

Competing financial interests

The authors declare no competing financial interests.

Author contributions

M.S.N. and F.H. designed the research, M.S.N., D.J.B., G.R., and S.P.-A. planned the analysis. E.H., E.D., S.V.G.Z.W., F.B., V.K., A.I., S.B.-M., Z.F., E.K., S.P.-A., and Z.K. collected data. M.S.N., K.S., N.K., G.R., H.J.F., J.T., E.M.H., D.I.D., and S.R supervised data collection. M.S.N. and S.P.-A. analyzed the data. M.S.N. drafted the manuscript and received comments from S.P.-A., N.K., K.S., D.J.B., H.J.F., E.M.H, and F.H.

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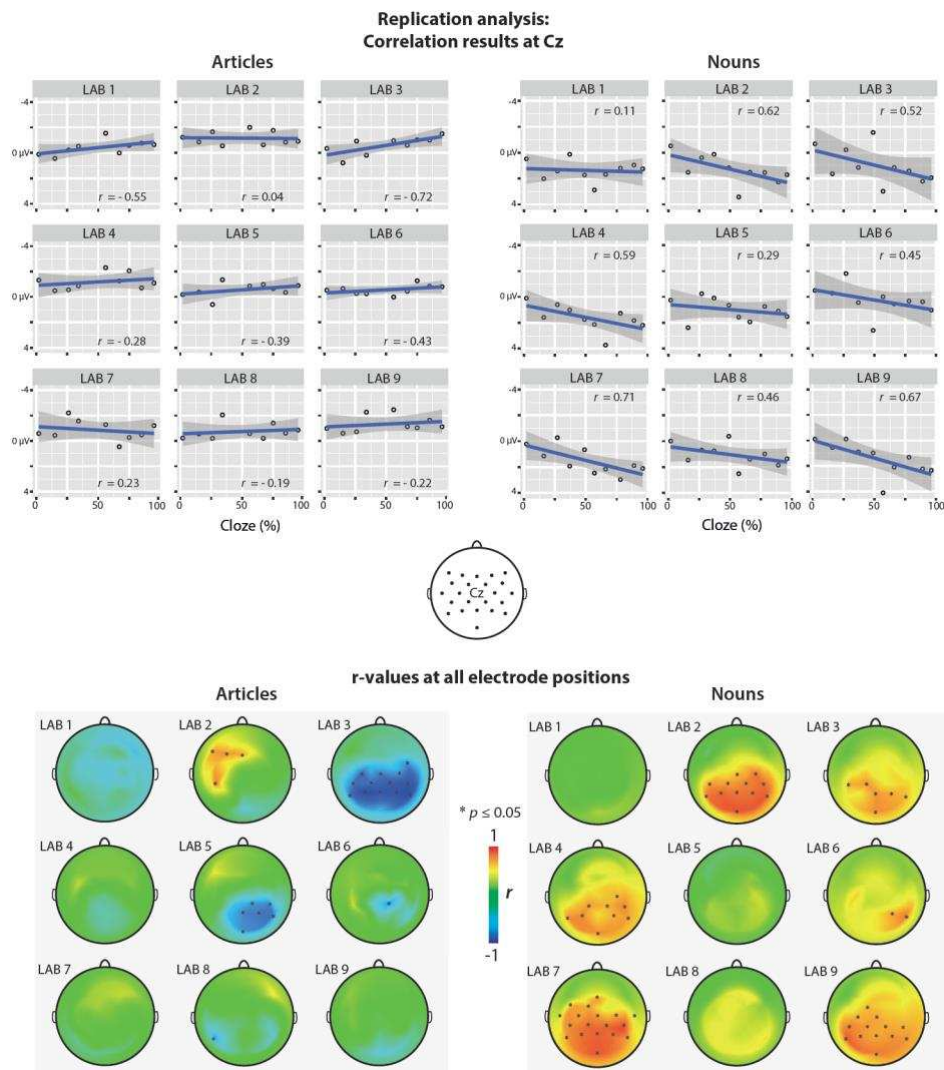
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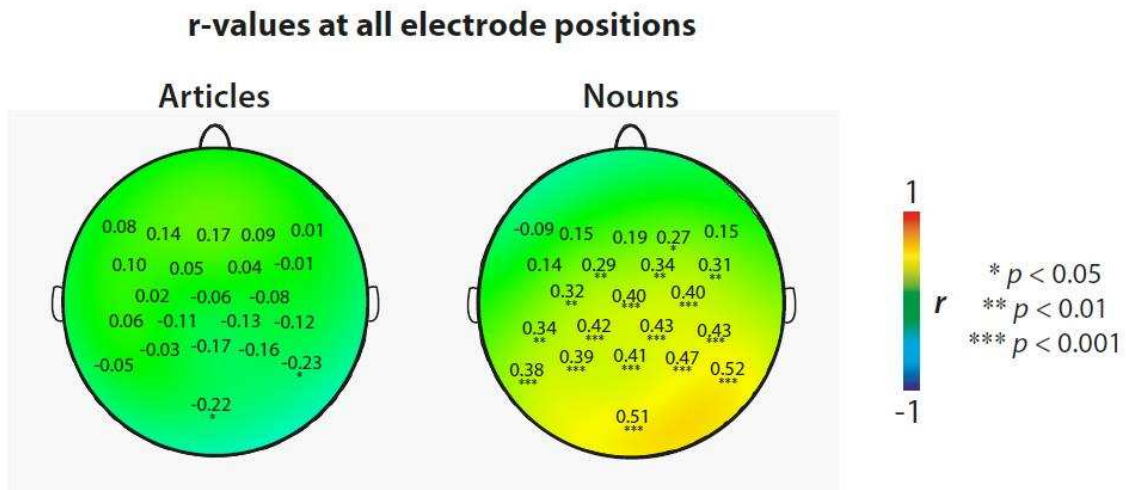
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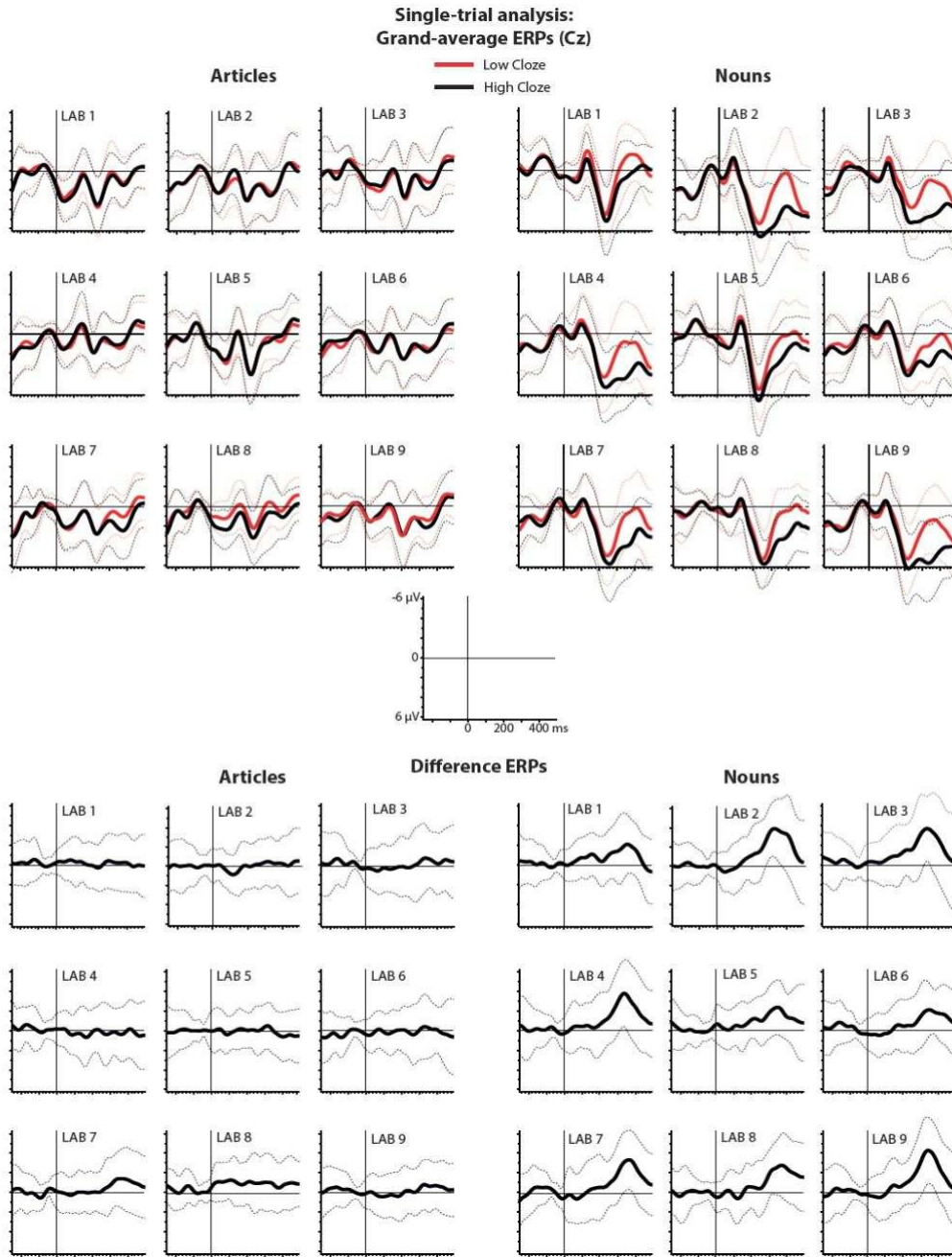
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791
 792 **Figure 1. Replication analysis.** Correlations between N400 amplitude and article/noun cloze
 793 probability per laboratory. N400 amplitude is the mean voltage in the 200-500 ms time
 794 window after word onset. A positive value corresponds to the canonical finding that N400
 795 amplitude became smaller (less negative—more positive) with increasing cloze probability.
 796 Here and in all further plots, negative voltages are plotted upwards. Upper graph: Scatter
 797 plots showing the correlation between cloze and N400 activity at electrode Cz, for each lab.
 798 The position of Cz and the other electrodes is displayed in the head plot in between the upper
 799 and lower graph. Lower graph: Scalp distribution of the r-values for each lab. Asterisks (*)
 800 indicate electrodes that showed a statistically significant correlation (two-tailed $p < 0.05$, not
 801 corrected for multiple comparisons). Exact r- and p-values for each laboratory and EEG
 802 channel are available on <https://osf.io/eyzaq>.
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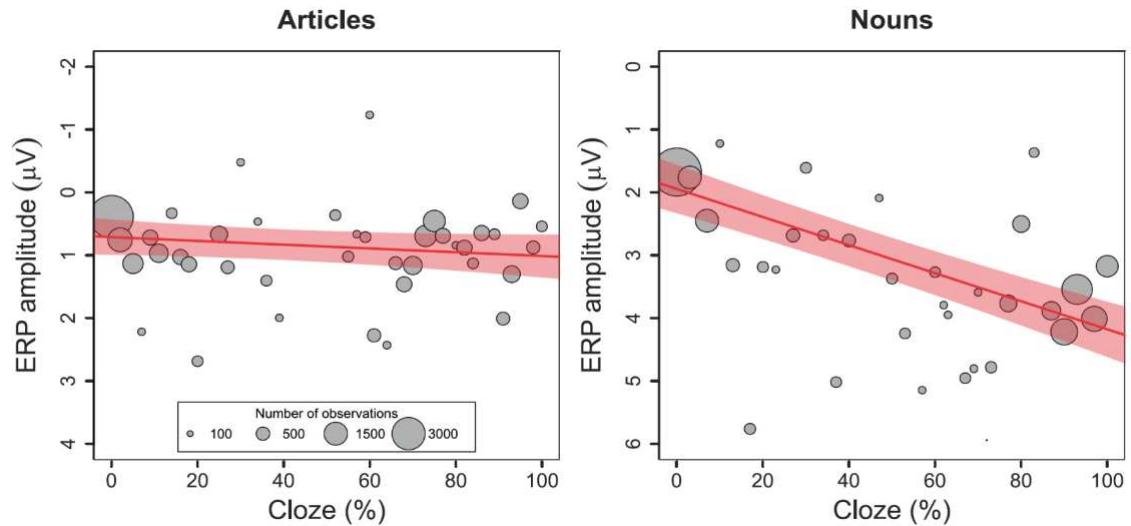


804
 805 **Figure 2. Replication analysis.** Scalp distribution and r-values at each channel based on data
 806 pooled from all laboratories, using a 500 ms baseline correction procedure as used by
 807 DeLong et al (2005). Asterisks (*) indicate electrodes that showed a statistically significant
 808 correlation (two-tailed, not corrected for multiple comparisons).

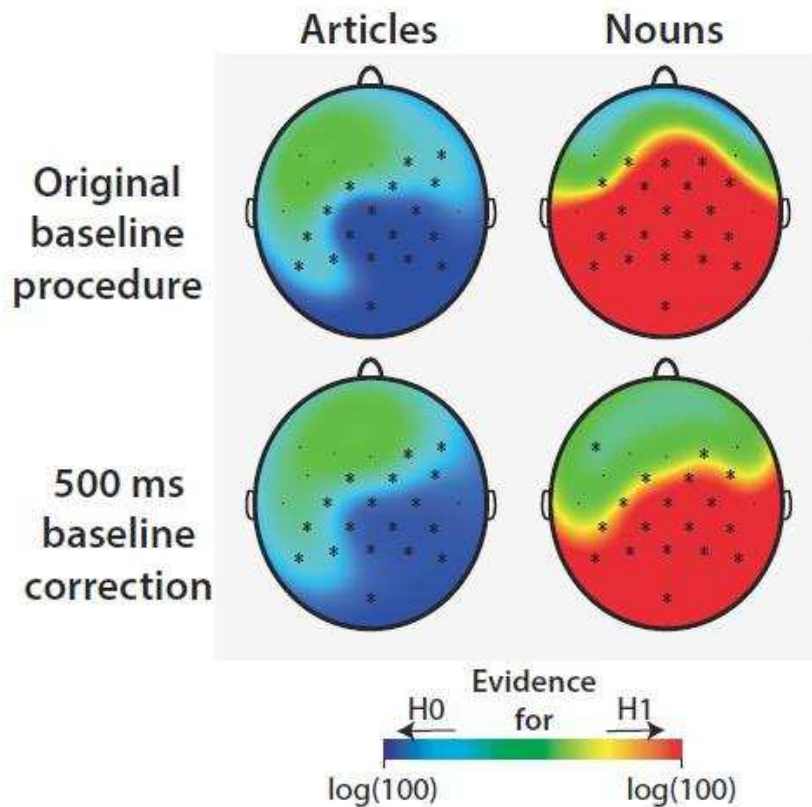


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810 **Figure 3. Single-trial analysis.** Grand-average ERPs elicited by relatively expected and
 811 unexpected words (cloze higher/lower than 50%) and the associated difference waveforms
 812 (low minus high cloze) at electrode Cz. Dotted lines indicate 1 standard deviation above or
 813 below the grand average.

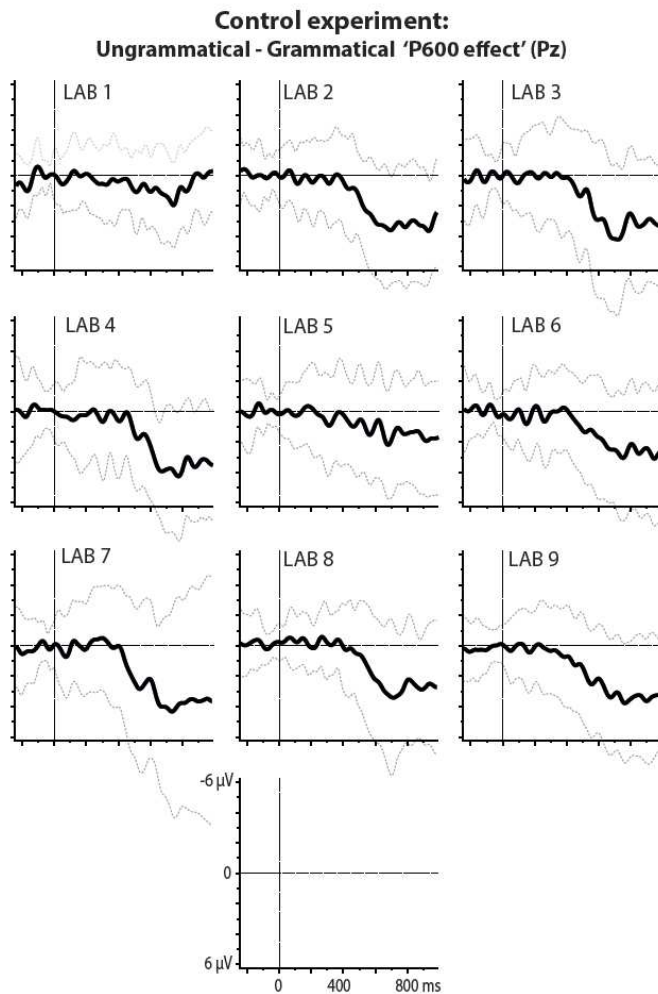


814 **Figure 4. Single-trial analysis.** Relationship between cloze and ERP amplitude for articles
 815 and nouns in the N400 spatiotemporal window, as illustrated by the mean ERP values per
 816 cloze value (number of observations reflected in circle size), along with the regression line
 817 and 95% confidence interval. A change in article cloze from 0 to 100 is associated with a
 818 change in amplitude of 0.296 μV (95% confidence interval: -.08 to .67). A change in noun-
 819 cloze from 0 to 100 is associated with a change in amplitude of 2.22 μV (95% confidence
 820 interval: 1.75 to 2.69). The data for these analyses was pooled across all 9 labs.
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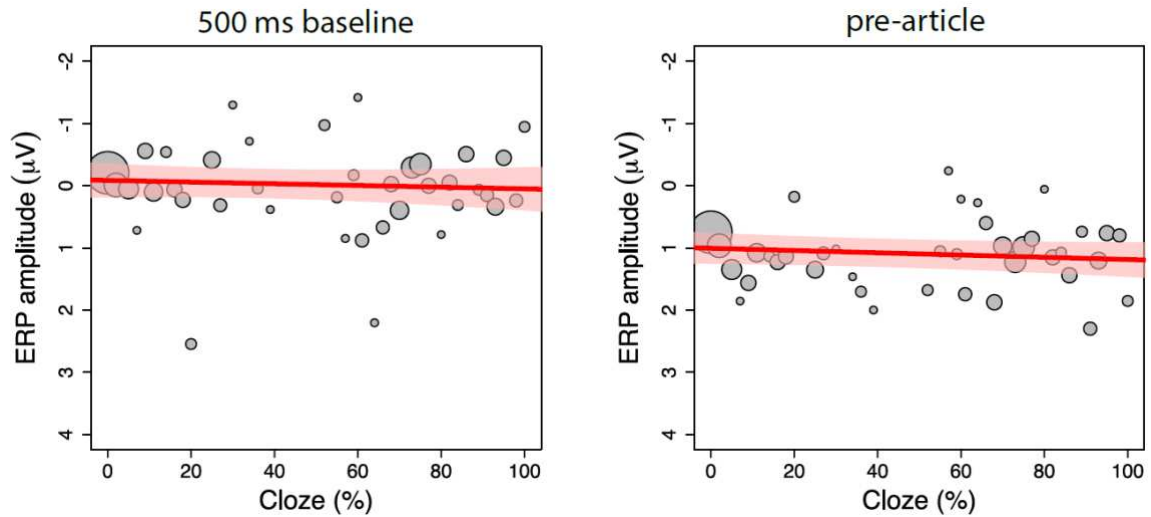
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823 **Figure 5.** Exploratory Bayes factor analysis associated with the replication analysis,
 824 quantifying the obtained evidence for the null hypothesis (H_0) that N400 is not impacted by
 825 cloze, or for the alternative hypothesis (H_1) that N400 is impacted by cloze with the direction
 826 and size of effect reported by DeLong et al. Scalp maps show the common logarithm of the
 827 replication Bayes factor for each electrode, capped at $\log(100)$ for presentation purposes.
 828 Electrodes that yielded at least moderate evidence for or against the null hypothesis (Bayes
 829 factor of ≥ 3) are marked by an asterisk. At posterior electrodes where DeLong et al. found
 830 their effects, our article data yielded strong to extremely strong evidence for the null
 831 hypothesis, whereas our noun data yielded extremely strong evidence for the alternative
 832 hypothesis (upper graphs). These results were obtained with the procedure described in
 833 DeLong et al. (no baseline correction), and with a 500 ms pre-word baseline correction
 834 (lower graphs), the procedure later described by DeLong and colleagues.



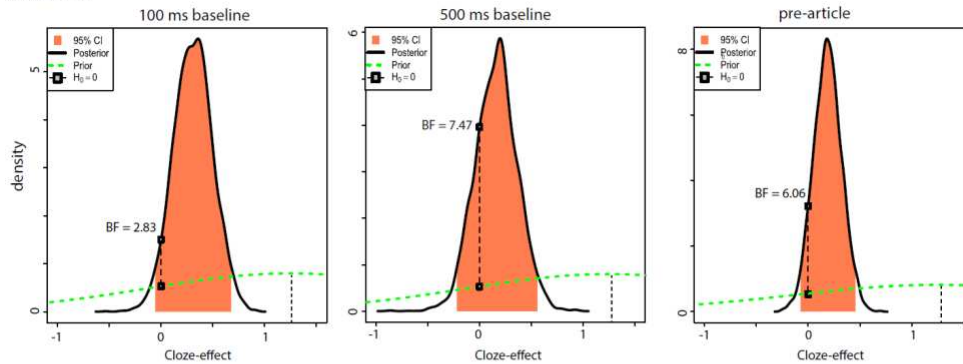
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836 **Figure 6. Control experiment.** P600 effects at electrode Pz per lab associated with flouting
 837 of the English a/an rule. Plotted ERPs show the grand-average difference waveform and
 838 standard deviation for ERPs elicited by ungrammatical expressions ('an kite') minus those
 839 elicited by grammatical expressions ('a kite').
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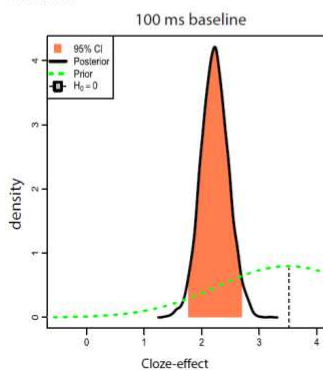


841 **Supplementary Figure 1. Exploratory single-trial analyses:** The relationship between cloze
 842 and ERP amplitude as illustrated by the mean ERP values per cloze value (number of
 843 observations reflected in circle size), along with the regression line and 95% confidence
 844 interval, from two exploratory analyses. We performed a test which used a longer baseline
 845 time windows (500 ms, left panel) to better control for pre-article voltage levels. This test
 846 reduced the initially observed effect of article-cloze, $\beta = .14$, CI [-.25, .53], $\chi^2(1) = 0.46$, $p =$
 847 $.50$). An analysis in the 500 to 100 ms time window before article-onset (right panel) revealed
 848 a non-significant effect of cloze that resembled the pattern observed after article-onset, $\beta =$
 849 $.16$, CI [-.07, .39], $\chi^2(1) = 1.82$, $p = .18$, shedding doubt on the conclusion that the observed
 850 results are due to the presentation of the articles.
 851

ARTICLES



NOUNS



852 **Supplementary Figure 2.** Results from exploratory Bayesian mixed-effects model analyses,
 853 represented by posterior distributions for the effect of cloze on ERP amplitudes in the N400
 854 window. The x-axis shows cloze effect sizes (i.e., changes in microvolts associated with an
 855 increase from 0% cloze probability to 100% cloze probability). The black line indicates the
 856 posterior distribution of effects; higher values of the posterior density at a given effect size
 857 indicate higher probability that this is the true effect size in the population. The peak of the
 858 posterior distribution roughly corresponds to the point estimate of the effect size (the
 859 regression coefficient) fitted from the Bayesian mixed effect model, i.e., the most likely value
 860 of the true effect size. The middle 95% of the posterior distribution, shaded in pink,
 861 corresponds to a two-tailed 95% credible interval for the effect size—i.e., an interval that we
 862 can be 95% confident contains the true effect. The green dotted line indicates the prior
 863 distribution (i.e., our expectation about where the true effect would lie before the data were
 864 collected). For the articles, this prior is centred on $1.25\mu\text{V}$, an approximation of the effect
 865 observed by DeLong and colleagues (2005), and for the nouns it is centred on $3.5\mu\text{V}$. The
 866 black connected dots illustrate the ratio between the posterior and prior distribution (i.e., the
 867 Bayes Factor) at the effect size of $0\mu\text{V}$; for example, a Bayes Factor of 4 suggests we can be
 868 4 times more certain that the true effect is zero after having conducted this experiment than
 869 before, or, in other words, that the data increased our confidence in the null effect of zero
 870 fourfold. We performed these analyses for each of the linear mixed-effects model analyses
 871 we performed. We note that in all the article-analyses, the posterior probability of the
 872 estimated effect being greater than zero is around 80 or 90%, although this is also true for the
 873 pre-stimulus variable, shedding doubt that the observed results are due to presentation of the
 874 articles. In none of our article-analyses did zero lie outside the obtained credible interval,
 875 whereas for the nouns, zero lay outside the credible interval. These results are consistent with
 876 a failure to replicate the size of the article-effect reported by DeLong et al. article-effect and
 877 successful replication of the noun-effect.