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INSURANCE LOSS COVERAGE AND DEMAND ELASTICITIES

BY MINGJIE HAO[†], ANGUS S. MACDONALD[‡], PRADIP TAPADAR[†] AND R. GUY THOMAS[†]

ABSTRACT

Restrictions on insurance risk classification may induce adverse selection, which is usually perceived as a bad outcome. We suggest a counter-argument to this perception in circumstances where modest levels of adverse selection lead to an increase in ‘loss coverage’, defined as expected losses compensated by insurance for the whole population. This happens if the shift in coverage towards higher risks under adverse selection more than offsets the fall in number of individuals insured. The possibility of this outcome depends on insurance demand elasticities for higher and lower risks. We state elasticity conditions which ensure that for any downward-sloping insurance demand functions, loss coverage when all risks are pooled at a common price is higher than under fully risk-differentiated prices. Empirical evidence suggests that these conditions may be realistic for some insurance markets.

KEYWORDS

Adverse selection; loss coverage; elasticity of demand; arc elasticity of demand.

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1. INTRODUCTION

Restrictions on insurance risk classification are common in life and health insurance markets. In the US, the Patient Protection and Affordable Care Act permits classification only by age, location, family size and smoking status; in the European Union, gender classification in insurance pricing has been banned; and many countries have restricted insurers’ use of genetic test results. Whilst such restrictions appear motivated by social objectives, they may also induce adverse selection, which is usually perceived as a bad outcome.

A simple version of the usual argument is as follows. If insurers are not permitted to charge risk-differentiated prices, they have to pool different risks at a common pooled

price. This pooled price is cheap for higher risks and expensive for lower risks; so more insurance is bought by higher risks, and less insurance is bought by lower risks. The equilibrium pooled price of insurance is higher than a population-weighted average of true risk premiums. Also, in most markets the number of higher risks is smaller than the number of lower risks, so the total number of risks insured falls.

However, the rise in equilibrium price under pooling reflects a shift in coverage towards higher risks. If the shift in coverage is large enough, it can more than offset the fall in numbers insured. In these circumstances, despite fewer risks being insured under pooling, expected losses compensated by insurance – a quantity we term ‘loss coverage’ – can be higher. We argue that where risk classification restrictions lead to higher ‘loss coverage’ – that is, more risk being voluntarily transferred and more losses being compensated – then from a social perspective, this should be seen as a good outcome from adverse selection.

The argument just given relies on a possibility, not a certainty. Whether the shift in coverage towards higher risks when risk classification is restricted is *in fact* large enough to offset the fall in numbers insured depends on the response of higher and lower risks to changes in the prices they face: the demand elasticities of higher and lower risks. This paper explores the relationship between insurance demand elasticities and loss coverage. Our main results state, for a sequence of increasingly general demand specifications and any number of risk-groups, the elasticity conditions which ensure that loss coverage will be higher when all risks are pooled at a common price than under fully risk-differentiated premiums.

This paper is related to previous literature as follows. Thomas (2008) first suggested loss coverage as a possible criterion for evaluating risk classification schemes, with illustrative examples generated using an exponential-power demand model suggested in De Jong and Ferris (2006). Hao et al. (2016a) gave a comprehensive mathematical analysis, but only for the case of iso-elastic demand with two risk-groups. Hao et al. (2016b), again focusing on iso-elastic demand, reconciled the concept of loss coverage to utilitarian concepts of social welfare found in related economic literature on risk classification, such as Hoy (2006), Einav and Finkelstein (2011) and Dionne and Rothschild (2014).

The rest of the paper is organised as follows. Section 2 presents a simple example which further clarifies and motivates the concept of loss coverage. Section 3 outlines our micro-founded model for insurance demand, where variations between individuals in utility functions lead to an aggregate proportional insurance demand between 0 and 1; this corresponds to the observable reality that not all individuals with the same probabilities of loss make the same insurance purchasing decisions. Section 4 sets up the model of risk classification, insurance market equilibrium and loss coverage. Section 5 establishes demand elasticity conditions for loss coverage to be higher under pooling than under risk-differentiated premiums, for three increasingly general specifications of insurance demand. Section 6 discusses how these conditions compare with empirical estimates of demand elasticities from various authors, and Section 7 gives conclusions.

2. SIMPLE EXAMPLE

The concept of loss coverage has been described elsewhere (Thomas (2008, 2009, 2017); Hao et al. (2016a,b)) but may remain unfamiliar to many readers, so this section gives a recap. In essence, the concept encapsulates an argument that the adverse selection induced by a ban on insurance risk classification can sometimes be beneficial for society as a whole.

The argument can be illustrated with a simple example, in the same spirit as dice-rolling examples to illustrate probability laws. Consider a population of just ten risks (say lives), and three alternative scenarios for risk classification: risk-differentiated premiums, pooled premiums (with some adverse selection), and pooled premiums (with severe adverse selection). We assume that all losses and insurance cover are for unit amounts (this simplifies the discussion, but it is not necessary). We assume the probability of loss is unaffected by the purchase of insurance (i.e. no moral hazard). The three scenarios are represented in the three panels of Figure 1.

In Figure 1, each H represents one higher risk and each L represents one lower risk. The population has the typical predominance of lower risks: a lower risk-group of eight risks each with probability of loss 0.01, and a higher risk-group of two risks each with probability of loss 0.04. In each scenario, the shaded cover above some of the H and L denote the risks covered by insurance.

In Scenario 1, risk-differentiated premiums (actuarially fair premiums) are charged. The proportion of each risk-group which buys insurance under these conditions is 50%, in line with industry statistics for e.g. life insurance.¹ The shading shows that a total of five risks are covered. Note that the equal areas of shading over four low risks and over one high risk symbolise equal expected losses.

The weighted average of the premiums paid is $(4 \times 0.01 + 1 \times 0.04)/5 = 0.016$. Since higher and lower risks are insured in the same proportions as they exist in the population, there is no adverse selection. The expected losses compensated by insurance for the whole population can be indexed by:

$$\text{Loss coverage} = \frac{\text{Expected compensated losses}}{\text{Expected population losses}} = \frac{4 \times 0.01 + 1 \times 0.04}{8 \times 0.01 + 2 \times 0.04} = 50\%. \quad (1)$$

In Scenario 2, in the middle panel of Figure 1, risk classification has been banned, and so insurers have to charge a common pooled premium to both higher and lower risks. Higher risks buy more insurance, and lower risks buy less. The weighted average premium

¹Some relevant industry statistics are as follows. The Life Insurance Market Research Organisation (LIMRA) states that 44% of US households have some individual life insurance (LIMRA (2013)). The US adult population (aged 18 years and over) at 1 July 2013 as estimated by the US Census Bureau was 244m; the American Council of Life Insurers states that 144m (59%) individual policies were in force in 2013 (ACLI, 2014, p72).

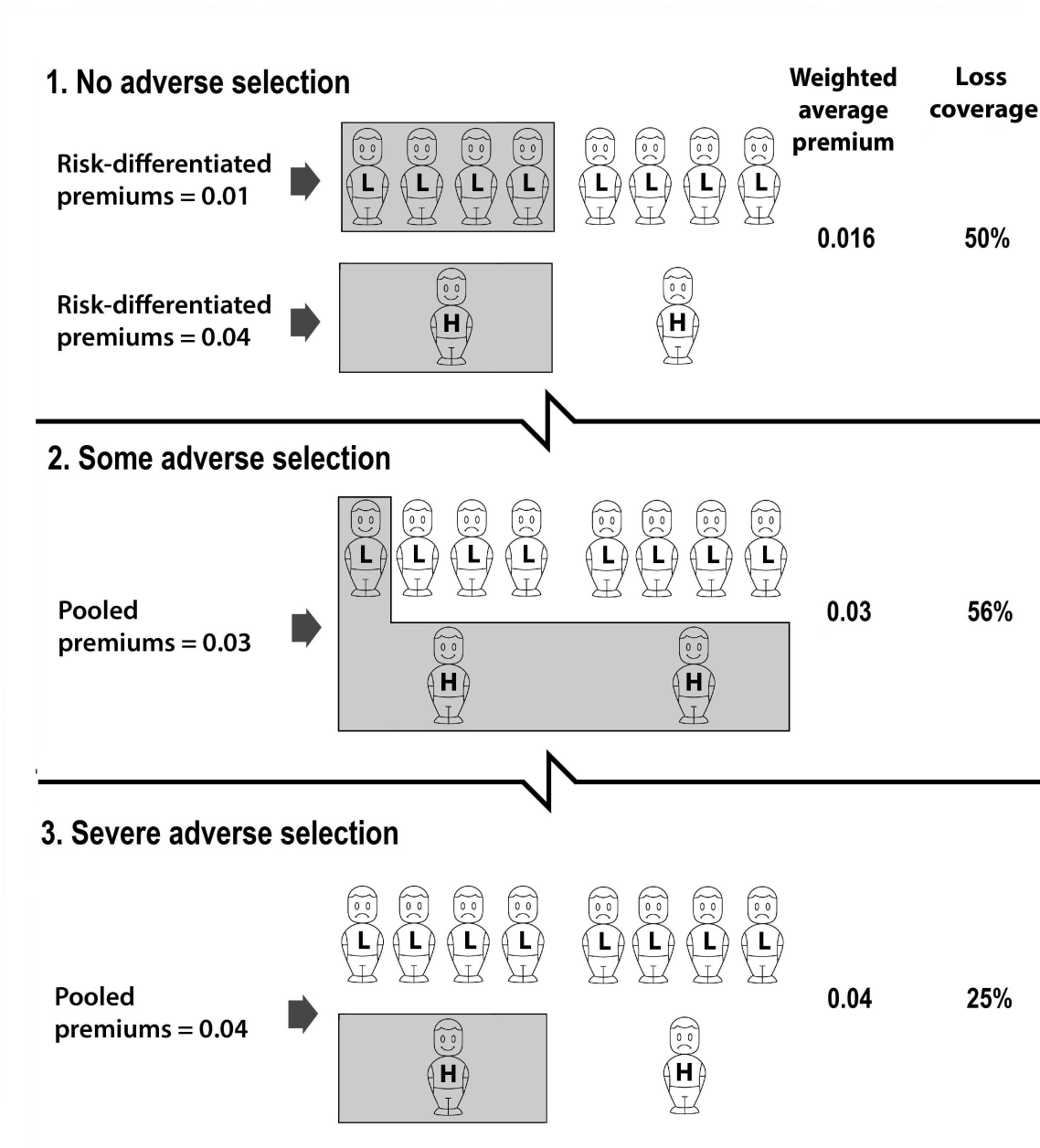


Figure 1: Three scenarios for risk classification (Adapted with permission from Thomas (2017)).

is $(1 \times 0.01 + 2 \times 0.04)/3 = 0.03$. The shading shows that three risks (compared with five previously) are now covered.

Note that the weighted average premium is higher in Scenario 2, and the number of risks insured is smaller. These are the essential features of adverse selection, which Scenario 2 accurately and completely represents. But there is a surprise: despite the adverse selection in the second scenario, the expected losses compensated by insurance for the whole population are now larger. Visually, this is symbolised by the *larger area of shading* in Scenario 2. Arithmetically, the loss coverage is:

$$\text{Loss coverage} = \frac{1 \times 0.01 + 2 \times 0.04}{8 \times 0.01 + 2 \times 0.04} = 56\%. \quad (2)$$

We argue that Scenario 2, with a higher expected fraction of the populations losses compensated by insurance – higher loss coverage – is superior from a social viewpoint to Scenario 1. The superiority of Scenario 2 arises not *despite* adverse selection, but *because* of adverse selection.

However, a ban on risk classification can also reduce loss coverage, if the adverse selection which the ban induces becomes too severe. This possibility is illustrated in Scenario 3, in the lower panel of Figure 1. Adverse selection has progressed to the point where only one higher risk, and no lower risks, buys insurance. Visually, the lower loss coverage is symbolised by the *smaller area* of shading in Scenario 3. Arithmetically, the loss coverage is:

$$\text{Loss coverage} = \frac{1 \times 0.04}{8 \times 0.01 + 2 \times 0.04} = 25\%. \quad (3)$$

These scenarios suggest that banning risk classification can increase loss coverage if it induces the ‘right amount’ of adverse selection (Scenario 2), but reduce loss coverage if it generates ‘too much’ adverse selection (Scenario 3). Which of Scenario 2 or Scenario 3 actually prevails depends on the demand elasticities of higher and lower risks. Hence it is of interest to explore the demand elasticity conditions under which loss coverage is increased by a ban on risk classification.

3. INSURANCE DEMAND

In the simple example, the possibility of an increase in loss coverage when risk classification was banned depended on the fact that not all higher-risk individuals chose to buy insurance at an actuarially fair premium. This corresponds to the reality of voluntary insurance markets (e.g. see the figures in Footnote 1). But it is not consistent with typical theories of insurance demand, which imply that all individuals will purchase full insurance if offered an actuarially fair premium (e.g. Mossin (1968)). This section presents a theory of insurance demand which accommodates the observable reality that not all individuals with the same probabilities of loss make the same insurance purchasing decision.

In Hao et al. (2016a), we assumed that some fixed proportion of individuals in each risk-group purchased insurance, consistent with the observation that in voluntary insurance markets, many individuals do not buy insurance. Thus, the model of insurance purchasing was at the level of the collective. In Hao et al. (2016b), we provided a micro-foundation for the collective model, based on personal utilities. We assumed that personal utilities within a risk-group are heterogeneous, so that different persons will make different decisions, when offered the same insurance premium. If the heterogeneity of utilities can be described in terms of a probability distribution, then the collective decision model of Hao et al. (2016a) can be recovered in terms of expected values.

This micro-foundation provides the ‘back-story’ for the proportional insurance demand functions which we use in Section 4 onwards. Because the ‘back-story’ is not the primary focus of this paper, the presentation here will be succinct. For a comprehensive account of the full probabilistic model underpinning this formulation, please refer to Hao et al. (2016b).

3.1 MICRO-FOUNDATIONS: HETEROGENEITY IN INDIVIDUAL RISK PREFERENCES

Consider an individual with initial wealth W , who risks losing an amount of L (with $L \leq W$) with probability μ . Suppose wealth preferences are governed by the utility function $U(w)$, which is increasing in wealth w , i.e. $U'(w) > 0$. (Individuals are typically also assumed to be risk-averse i.e. $U''(w) < 0$, but our theory of insurance demand does *not* require that *all* individuals are risk-averse.)

Suppose that the individual is offered insurance providing full cover against the potential loss amount L at premium π per unit of loss, i.e. for a payment of πL . (In this paper we do not consider partial cover.) She will choose to buy insurance if π is low enough to satisfy:

$$U(W - \pi L) > (1 - \mu)U(W) + \mu U(W - L). \quad (4)$$

Now consider a group of individuals, all with the same initial wealth W and the same potential loss amount L (both fixed and with $L \leq W$). Then in the above model, all individuals with the same probability of loss μ make the same purchasing decision, depending only on whether or not the offered premium π is small enough to satisfy Equation 4. In reality, it appears implausible that all individuals with the same probability of loss make the same purchasing decision (e.g. for life insurance, see the figures in footnote 1).

A possible explanation for this apparent inconsistency is that risk preferences are heterogeneous (see e.g. Finkelstein and McGarry (2006); Cutler et al. (2008)).² That

²Two other possible explanations for different purchasing decisions by individuals with the same probabilities of loss who are offered the same premium are variations across individuals in initial wealth W , and variations across individuals in the potential loss amount L . We do not consider these explanations in this paper.

is, in a risk-group in which all individuals have the same risk μ , individuals may have different utility functions. Suppose for simplicity that utility functions belong to a family parameterized by a positive real number γ . So a particular individual's utility function can be denoted by $U_\gamma(w)$.

Further suppose that, within the risk-group, γ is sampled randomly from some random variable Γ with distribution function $F_\Gamma(\gamma)$. So, a particular individual's utility function, $U_\gamma(w)$, is a random quantity. Any numerical quantity involving an individual's utility function is therefore a random variable, the randomness being inherited from the distribution $F_\Gamma(\gamma)$.

Now, a single individual in the risk-group will choose to buy insurance at premium π if and only if Equation 4 is satisfied for their particular utility function $U_\gamma(w)$:

$$U_\gamma(W - \pi L) > (1 - \mu)U_\gamma(W) + \mu U_\gamma(W - L). \quad (5)$$

Note that this behaviour is completely deterministic, assuming that individuals know their own preferences.

Since decisions based on Equation 5 do not depend on the origin and scale of a utility function, we will find it convenient to assume that all individuals have the same utility at the 'end points' $W - L$ and W . The following standardisation is convenient: $U_\gamma(W) = 1$ and $U_\gamma(W - L) = 0$. Equation (5) then becomes:

$$U_\gamma(W - \pi L) > (1 - \mu). \quad (6)$$

The utility at the fixed wealth $(W - \pi L)$ is a random variable, that we denote by $U_\Gamma(W - \pi L)$, with a distribution function induced by the random variable Γ .

We now make the key assumption, that while individuals know their own utility function, this is unobservable to the insurer. Insurers can observe the probability of loss μ , and naturally know the offered premium π , but within the risk-group the insurer can at most observe the proportion of individuals who choose to buy insurance. We call this a (proportional) demand function and define it as:

$$d(\pi) = \text{P} [U_\Gamma(W - \pi L) > (1 - \mu)] \quad (7)$$

where P denotes a probability.

Note that each individual's decision is completely deterministic, given their knowledge of their own utility function; but to the insurer it appears stochastic, given what the insurer knows. In respect of any individual chosen randomly, define the function Q to be $Q = 1$ if they buy insurance or $Q = 0$ if they do not. To the individual concerned, Q is a deterministic function. To the insurer, Q is a Bernoulli random variable with parameter $d(\pi)$.

As a concrete example, Result A.1 of Appendix A shows the iso-elastic demand function obtained in the case of power utility functions parameterized by a random variable Γ with a specified distribution.

A notable feature of this model is that demand when an actuarially fair premium is charged can be less than 1. By permitting some individuals to be ‘risk lovers’, the model can generate a demand function consistent with the observation that not everyone buys insurance (e.g. the figures in footnote 1). Although ‘risk-loving’ or ‘risk-seeking’ are the usual stylised descriptions, it suffices for our model that some individuals, *for whatever reason*, do not purchase insurance when offered an actuarially fair premium. In many cases ‘risk-neglecting’ might be a more realistic description than ‘risk-seeking’.

Another possible explanation of some individuals’ non-purchasing may be that an actuarially fair premium is, in practice, not offered, i.e. premium loadings for expenses and profit. However, heterogeneity in risk aversion is a more flexible explanation, in that it can explain why different individuals with the same probabilities of loss and offered the same premiums may make different decisions.

3.2 AGGREGATES: PROPORTIONAL DEMAND AND DEMAND ELASTICITIES

The micro-foundations described above provide a possible mechanism by which aggregate demand for insurance is generated by the unobserved utility functions of individuals. In the remainder of this paper, we work directly with the aggregate quantity, the proportional demand $d(\pi)$ for insurance ($0 \leq d(\pi) \leq 1$).

We now define a related concept, the (point price) elasticity of insurance demand, as follows:

$$\epsilon(\pi) = -\frac{\partial \log d(\pi)}{\partial \log \pi}, \quad \text{so consequently:} \quad (8)$$

$$d(\pi) = \tau \exp \left[- \int_{\mu}^{\pi} \epsilon(s) d \log s \right], \quad (9)$$

where the parameter $\tau = d(\mu)$ is the *fair-premium demand* for insurance. The expression in Equation 9 has the benefit that we do not have to impose any regularity conditions, like continuity and differentiability, on the demand function $d(\pi)$. We require only that the demand elasticity $\epsilon(\pi)$ in Equation 8 is non-negative, so that the demand $d(\pi)$ in Equation 9 is non-increasing in premium π .

4. EQUILIBRIUM AND LOSS COVERAGE FOR TWO OR MORE RISK-GROUPS

4.1 FRAMEWORK FOR INSURANCE RISK CLASSIFICATION

In Section 3, we suggested a model for insurance demand from a group of individuals who all have the same probability of loss but who may have different risk preferences. We

call such a group of individuals a *risk-group*. In this section, we extend the model to a population consisting of several risk-groups with different loss probabilities.

Suppose a population consists of n distinct risk-groups with probabilities of loss given by $\mu_1, \mu_2, \dots, \mu_n$. For convenience, we assume that $0 < \mu_1 < \mu_2 < \dots < \mu_n < 1$.

Suppose the proportion of the population belonging to risk-group i is p_i , for $i = 1, 2, \dots, n$. If we choose an individual at random from the population, their probability of loss is a random variable, which we denote by μ , and its distribution is given by $P[\mu = \mu_i] = p_i$ for $i = 1, 2, \dots, n$.

Suppose insurers charge premiums (per unit of loss) $\pi_1, \pi_2, \dots, \pi_n$ for the risk-groups $i = 1, 2, \dots, n$, respectively. Based on the model of Section 3, the demand for insurance within risk-group i is denoted by $d_i(\pi_i)$, where $0 \leq d_i(\pi_i) \leq 1$ and $d_i(\pi_i)$ is non-increasing in π_i .

Let the insurance purchasing decision of an individual chosen at random from the whole population be represented by the indicator random variable Q , taking the value of 1 if insurance is purchased; and 0 otherwise. Within risk-group i , this has expectation:

$$E[Q \mid \mu = \mu_i] = P[Q = 1 \mid \mu = \mu_i] = d_i(\pi_i). \quad (10)$$

and the unconditional expectation is:

$$E[Q] = \sum_{i=1}^n E[Q \mid \mu = \mu_i] P[\mu = \mu_i] = \sum_{i=1}^n d_i(\pi_i) p_i. \quad (11)$$

which represents the expected fraction of the whole population who buy insurance, ie. the expected demand for insurance.

Now suppose that the occurrence of a loss event for an individual chosen at random from the whole population is represented by the indicator random variable, X , taking the value of 1 if a loss event occurs; and 0 otherwise. Within risk-group i , this has expectation:

$$E[X \mid \mu = \mu_i] = P[X = 1 \mid \mu = \mu_i] = \mu_i. \quad (12)$$

and the unconditional expectation is:

$$E[X] = \sum_{i=1}^n E[X \mid \mu = \mu_i] P[\mu = \mu_i] = \sum_{i=1}^n \mu_i p_i \quad (13)$$

which represents the expected fraction of the population who suffer a loss.

Conditional on $\mu = \mu_i$, we assume that Q and X are independent. This ensures that there is no moral hazard; although the level of risk may influence the decision to buy insurance, mediated by $d_i(\pi_i)$, insured individuals in any risk-group have the same probability of loss as uninsured individuals.

Now recall that the loss, if it occurs, is always of fixed amount L . The insurance claim in respect of an individual chosen at random from the whole population is the cover indicator (Q) times the loss indicator (X) times the fixed loss amount (L). Its expectation is:

$$\begin{aligned}
\mathbb{E}[QXL] &= \sum_{i=1}^n \mathbb{E}[QXL \mid \mu = \mu_i] \mathbb{P}[\mu = \mu_i], \\
&= L \sum_{i=1}^n \mathbb{E}[Q \mid \mu = \mu_i] \mathbb{E}[X \mid \mu = \mu_i] \mathbb{P}[\mu = \mu_i], \\
&= L \sum_{i=1}^n d_i(\pi_i) \mu_i p_i.
\end{aligned} \tag{14}$$

Next, define Π to be the premium which would be chargeable to an individual chosen at random from the population, if that individual purchased insurance. Then Π is a random variable, taking values $\pi_1, \pi_2, \dots, \pi_n$ for individuals from risk-groups $i = 1, 2, \dots, n$, respectively. The premium income in respect of an individual chosen at random from the whole population is the cover indicator (Q) times the premium (Π) times the fixed loss amount (L). Its expectation is:

$$\begin{aligned}
\mathbb{E}[Q\Pi L] &= \sum_{i=1}^n \mathbb{E}[Q\Pi L \mid \mu = \mu_i] \mathbb{P}[\mu = \mu_i], \\
&= L \sum_{i=1}^n \mathbb{E}[Q \mid \mu = \mu_i] \mathbb{E}[\Pi \mid \mu = \mu_i] \mathbb{P}[\mu = \mu_i], \\
&= L \sum_{i=1}^n d_i(\pi_i) \pi_i p_i
\end{aligned} \tag{15}$$

We call any vector of premiums $(\pi_1, \pi_2, \dots, \pi_n)$ charged by insurers a risk classification regime, and denote it by underlined characters such as $\underline{\pi}$. The insurer's expected profit under risk classification regime $\underline{\pi}$, which we denote by $\rho(\underline{\pi})$ is then:

$$\rho(\underline{\pi}) = \mathbb{E}[Q\Pi L] - \mathbb{E}[QXL] = L \sum_{i=1}^n d_i(\pi_i) \pi_i p_i - L \sum_{i=1}^n d_i(\pi_i) \mu_i p_i. \tag{16}$$

We assume that competition between insurers leads to zero profits in equilibrium, that is $\rho(\underline{\pi}) = 0$. Dividing both sides by the fixed loss amount L which appears in Equation 16, we can write this as

$$\sum_{i=1}^n d_i(\pi_i) p_i(\pi_i - \mu_i) = 0. \quad (17)$$

which we refer to as the *equilibrium condition*.

We also assume that competition drives all insurers to classify risks to the maximum extent which regulation permits. Depending on applicable regulation, two polar cases of risk classification regime are as follows:

- Full risk classification, under which $\pi_i = \mu_i$. Here, the insurer uses the maximum possible degree of underwriting. We denote this regime by $\underline{\pi} = \underline{\mu} = (\mu_1, \mu_2, \dots, \mu_n)$.
- No risk classification, or risk pooling, under which all $\pi_i = \pi$, a constant. Here, the insurer uses the minimum possible degree of underwriting.

By considering the insurer's profit under risk pooling with $\pi = \mu_1$ and $\pi = \mu_n$, it is clear that there must be at least one risk pooling regime which satisfies the equilibrium condition.³

4.2 LOSS COVERAGE

We define loss coverage under a risk classification $\underline{\pi}$ that satisfies the equilibrium condition as the expected losses across the whole population that are compensated by insurance, i.e. $E[QXL]$ as defined in Equation 14. That is:

$$\text{Loss coverage: } LC(\underline{\pi}) = E[QXL], \quad (18)$$

where $\rho(\underline{\pi}) = 0$.

For comparison purposes we use loss coverage under full risk classification regime as a reference level, and hence define the *loss coverage ratio* as follows:

$$\text{Loss coverage ratio: } C = \frac{LC(\underline{\pi})}{LC(\underline{\mu})}. \quad (19)$$

Note that the fixed loss amount L cancels in the numerator and denominator of Equation 19, and so we shall subsequently omit it.

5. IMPACT OF DEMAND ELASTICITIES ON LOSS COVERAGE

³In general, there may be more than one, and the results in this paper allow for this possibility. But such multiple solutions do not arise for typical demand functions and elasticities. For more details see Hao et al. (2016a) and Appendix B of Thomas (2017). Between the polar cases of full risk classification and pooling, many schemes of 'partial risk classification' are possible; for some general discussion see Chapter 6 of Thomas (2017).

In this section we investigate the impact of demand elasticities for higher and lower risks on loss coverage ratio. For ease of exposition we start with iso-elastic demand where all risk-groups have the same elasticity λ , and use this as a stepping stone to prove successively more general results.

5.1 SAME ISO-ELASTIC DEMAND ELASTICITY FOR ALL RISK-GROUPS

Suppose the elasticity of demand is the same constant for all individuals, irrespective of their risks and of the premium charged, i.e. insurance demand is iso-elastic. So:

$$\epsilon_i(\pi) = \lambda \text{ (a positive constant), for } i = 1, 2, \dots, n. \quad (20)$$

The resulting demand function found from Equation 9 is:

$$d_i(\pi) = \tau_i \left(\frac{\mu_i}{\pi} \right)^\lambda, \quad i = 1, 2, \dots, n. \quad (21)$$

where the parameter $\tau_i = d_i(\mu_i)$ is the *fair-premium demand* for insurance, i.e. the demand when an actuarially fair premium is charged. We assume $\tau_i < 1$ for all risk-groups.

We note that for any risk-group, demand must logically be subject to a cap of 1 (i.e. once the premium charged is low enough so that all members of a risk-group buy insurance, further reductions in premium do not lead to further increases in demand). Our subsequent analysis in this paper assumes that demand is less than 1 for all risk-groups at the pooled premium. It would be prudent to verify this in any applications, particularly for higher risk-groups.

Based on iso-elastic demand as in Equation 21, the equilibrium condition in Equation 17 under risk pooling with premium π_0 gives:

$$\sum_{i=1}^n p_i \tau_i \left(\frac{\mu_i}{\pi_0} \right)^\lambda (\pi_0 - \mu_i) = 0, \text{ or, equivalently:} \quad (22)$$

$$\sum_{i=1}^n \alpha_i \left(\frac{\mu_i}{\pi_0} \right)^\lambda = \sum_{i=1}^n \alpha_i \left(\frac{\mu_i}{\pi_0} \right)^{\lambda+1}, \text{ where } \alpha_i = \frac{p_i \tau_i}{\sum_{i=1}^n p_i \tau_i} \text{ for } i = 1, 2, \dots, n, \quad (23)$$

and the unique pooled equilibrium premium, π_0 , is given by:

$$\pi_0 = \frac{\sum_{i=1}^n \alpha_i \mu_i^{\lambda+1}}{\sum_{i=1}^n \alpha_i \mu_i^\lambda} \quad (24)$$

To facilitate mathematical proofs, it is helpful to re-frame Equation 23 as follows. Consider a random variable, V , taking values $v_i = \frac{\mu_i}{\pi_0}$ with probabilities α_i for $i = 1, 2, \dots, n$. Then Equation 23 says that under equilibrium, the random variable V satisfies:

$$\text{E} [V^\lambda] = \text{E} [V^{\lambda+1}]. \quad (25)$$

The loss coverage ratio, as defined in Equation 19, comparing loss coverage under pooled premiums to that under risk-differentiated premiums, can be re-framed as:

$$C = \frac{LC(\pi_0)}{LC(\underline{\mu})} = \frac{\sum_{i=1}^n \alpha_i \left(\frac{\mu_i}{\pi_0}\right)^\lambda \mu_i}{\sum_{i=1}^n \alpha_i \mu_i} = \frac{\sum_{i=1}^n \alpha_i \left(\frac{\mu_i}{\pi_0}\right)^{\lambda+1}}{\sum_{i=1}^n \alpha_i \left(\frac{\mu_i}{\pi_0}\right)} = \frac{E[V^{\lambda+1}]}{E[V]} = \frac{E[V^\lambda]}{E[V]}. \quad (26)$$

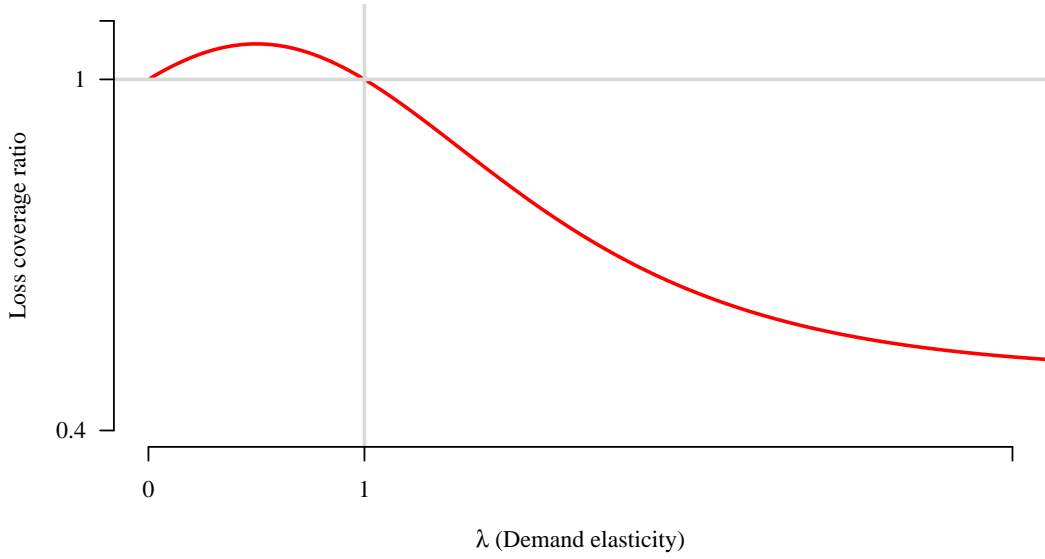


Figure 2: Loss coverage ratio as a function of demand elasticity for iso-elastic demand.

Figure 2 shows the plot of loss coverage ratio as a function of demand elasticity λ , for the risks $(\mu_1, \mu_2) = (0.01, 0.04)$ and $(\alpha_1, \alpha_2) = (0.8, 0.2)$. We can see that loss coverage under pooling is higher than under risk-differentiated premiums if demand elasticity is less than 1, and vice versa. The pattern shown in Figure 2 is formally stated in the following result:

Result 5.1. *Suppose there are n risk-groups with risks $\mu_1 < \mu_2 < \dots < \mu_n$ and the same iso-elastic demand elasticity λ . Then $\lambda \lesseqgtr 1 \Rightarrow C \gtrless 1$.*

This result generalises a result proved for only two risk-groups in Hao et al. (2016a). A proof is given in Appendix B.

5.2 DIFFERENT ISO-ELASTIC DEMAND ELASTICITIES FOR DIFFERENT RISK-GROUPS

The assumption of constant iso-elastic demand elasticity for all individuals, although mathematically tractable, can be criticised as unrealistic. For most goods and services, we expect demand elasticity to rise with price, because of the income effect on demand: at higher prices, the good forms a larger part of the consumer's total budget constraint, and so the effect of a small percentage change in its price might be larger. For insurance this suggests that demand elasticity for higher risks (who are typically charged higher prices) might be higher. So we next consider iso-elastic demand functions with different demand elasticities for different risk-groups, i.e.:

$$\epsilon_i(\pi) = \lambda_i \text{ for } i = 1, 2, \dots, n; \quad (27)$$

where higher risk-groups are likely to have higher demand elasticities. Under this formulation, the equilibrium condition under risk pooling with premium π_0 gives:

$$\sum_{i=1}^n p_i \tau_i \left(\frac{\mu_i}{\pi_0} \right)^{\lambda_i} (\pi_0 - \mu_i) = 0, \text{ or, equivalently:} \quad (28)$$

$$\sum_{i=1}^n \alpha_i \left(\frac{\mu_i}{\pi_0} \right)^{\lambda_i} = \sum_{i=1}^n \alpha_i \left(\frac{\mu_i}{\pi_0} \right)^{\lambda_i + 1}. \quad (29)$$

As in Section 5.1, we define a random variable V taking values $v_i = \frac{\mu_i}{\pi_0}$ with probabilities α_i for $i = 1, 2, \dots, n$. Now, define a function $f(v)$, such that:

$$f(v_i) = \lambda_i, \text{ for } i = 1, 2, \dots, n. \quad (30)$$

Then the equilibrium condition, in Equation 29, can be re-framed as:

$$\mathbb{E} [V^{f(V)}] = \mathbb{E} [V^{f(V)+1}]. \quad (31)$$

Under this setting, we have the following result:

Result 5.2. *Suppose there are n risk-groups with risks $\mu_1 < \mu_2 < \dots < \mu_n$ with iso-elastic demand elasticities $\lambda_1, \lambda_2, \dots, \lambda_n$ respectively. Define $\lambda_{lo} = \max_{v \leq 1} f(v)$ and $\lambda_{hi} = \min_{v > 1} f(v)$. Then $\lambda_{lo} < 1$ and $\lambda_{hi} \geq \lambda_{lo} \Rightarrow C \geq 1$.*

To understand this result, note that under pooled equilibrium, λ_{lo} is the maximum of the demand elasticities of all those *lower* risk-groups who pay more premium, π_0 , than their actuarially fair risk μ . So, $\lambda_{lo} < 1$ signifies that, for these *lower* risk-groups, their iso-elastic demand elasticities should be less than 1, which is consistent with empirical evidence in many markets (as discussed in Section 6 of this paper).

On the other hand, λ_{hi} , is the minimum of the demand elasticities of all those *higher* risk-groups who pay less premium, π_0 , than their actuarially fair risk, μ . The interpretation of the second condition, $\lambda_{hi} \geq \lambda_{lo}$, is that for these *higher* risk-groups, the demand

elasticities should be larger than those of all the *lower* risk-groups, which is consistent with what we expect from the income effect on demand.

In summary, as long as the iso-elastic demand elasticities of the different risk-groups satisfy the two conditions: $\lambda_{lo} < 1$ and $\lambda_{hi} \geq \lambda_{lo}$, the loss coverage under pooling is higher than under full risk classification.

The following special cases of Result 5.2 are worth noting:

1. If the iso-elastic demand elasticities are the same for all risk-groups, i.e. $\lambda_i = \lambda$ for $i = 1, 2, \dots, n$, then by definition $\lambda_{lo} = \lambda_{hi} = \lambda$, and so $\lambda_{lo} = \lambda < 1$ gives $C \geq 1$, which corresponds to Result 5.1.
2. For the special case of: $0 < \lambda_1 \leq \lambda_2 \leq \dots \leq \lambda_n < 1$, the two conditions, $\lambda_{lo} < 1$ and $\lambda_{hi} \geq \lambda_{lo}$, are trivially satisfied and hence in this case: $C \geq 1$.
3. For the case of two risk-groups, $\lambda_{lo} = \lambda_1$ and $\lambda_{hi} = \lambda_2$, and the conditions on the demand elasticities translate to $\lambda_1 < 1$ and $\lambda_2 \geq \lambda_1$.

The two risk-groups case is illustrated in Figure 3, where $(\mu_1, \mu_2) = (0.1, 0.4)$ and $(\alpha_1, \alpha_2) = (0.8, 0.2)$. The two axes represents λ_1 and λ_2 . The figure demarcates the region of $C > 1$ (shaded green) from the region of $C < 1$ (shaded pink) by the boundary curve $C = 1$ (in red).

The conditions on λ_1 and λ_2 say that in the region above the $\lambda_1 = \lambda_2$ diagonal and $\lambda_1 < 1$, demarcated by the green dashed borders, loss coverage under pooling is always higher than that under full risk classification. This is true irrespective of the relative sizes and relative risks of the higher and lower risk populations.

Figure 3 highlights another important point: that Result 5.2 focuses only on loss coverage inside the region demarcated by green dashes. Outside this region, loss coverage under pooling can be higher or lower than under full risk-classification (higher in the green segments to the right of the dashed green lines; lower throughout the red area towards the right). The position of the $C = 1$ curve which demarcates the red and green areas changes slightly with relative population sizes and relative risks. The region demarcated by green dashes is the only region for which we can obtain a universal result (i.e. one which holds independent of relative sizes and risks of higher and lower risk populations). Fortunately, empirical evidence and economic rationale imply that realistic values of (λ_1, λ_2) may often fall within this region.

A proof of Result 5.2 is given in Appendix C.

5.3 GENERAL DEMAND ELASTICITY FUNCTIONS

So far, we have only considered constant demand elasticities (as a function of premium), either for all individuals in the population, or for all individuals belonging to a particular risk-group. However, it can be argued that demand elasticities should actually

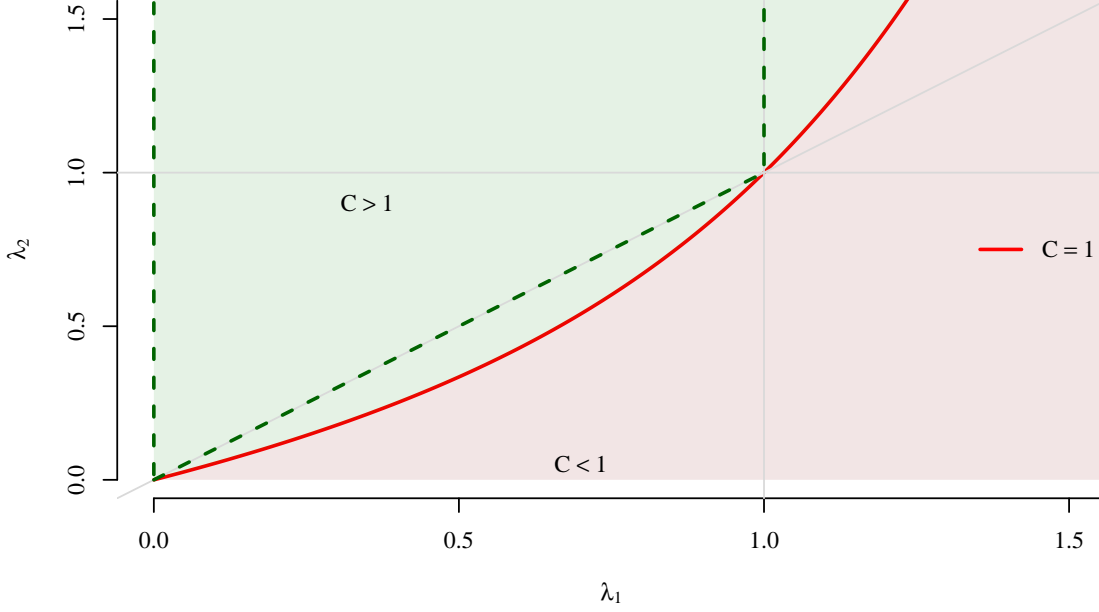


Figure 3: Curve demarcating the regions where loss coverage ratio is greater than and less than 1 when $(\mu_1, \mu_2) = (0.01, 0.04)$ and $(\alpha_1, \alpha_2) = (0.8, 0.2)$.

be increasing functions of premium (instead of being a constant), to reflect the income effect on demand; the argument being that at higher prices, insurance forms a larger part of the consumer's total budget constraint. In this section, we generalise our analysis to allow for different demand elasticity functions, $\epsilon_i(\pi)$, for different risk-groups $i = 1, 2, \dots, n$.

Using Equation 9, the proportional demand for insurance, for risk-group i , is:

$$d_i(\pi) = \tau_i \exp \left[- \int_{\mu_i}^{\pi} \epsilon_i(s) d \log s \right], \text{ for } i = 1, 2, \dots, n. \quad (32)$$

Under this formulation, the equilibrium condition under risk pooling with premium π_0 gives:

$$\sum_{i=1}^n p_i \tau_i \exp \left[- \int_{\mu_i}^{\pi_0} \epsilon_i(s) d \log s \right] (\pi_0 - \mu_i) = 0, \text{ or, equivalently:} \quad (33)$$

$$\sum_{i=1}^n \alpha_i \exp \left[\int_{\pi_0}^{\mu_i} \epsilon_i(s) d \log s \right] (\pi_0 - \mu_i) = 0; \quad (34)$$

in which the term:

$$\int_{\pi_0}^{\mu_i} \epsilon_i(s) d \log s, \quad (35)$$

can be interpreted using the concept of *arc elasticity of demand*, denoted by $\eta_i(a, b)$ and defined in Vazquez (1995) as follows:

$$\eta_i(a, b) = \frac{\int_a^b \epsilon_i(s) d \log s}{\int_a^b d \log s}. \quad (36)$$

Arc elasticity, $\eta_i(a, b)$, can be interpreted as the *average* value of (point) elasticity of demand for risk-group i , $\epsilon_i(s)$, over the price logarithmic arc from price a to price b . So in our case, we can define:

$$\lambda_i = \eta_i(\pi_0, \mu_i) = \frac{\int_{\pi_0}^{\mu_i} \epsilon_i(s) d \log s}{\int_{\pi_0}^{\mu_i} d \log s}, \text{ for } i = 1, 2, \dots, n. \quad (37)$$

Equation 34 can then be rewritten using arc elasticities as follows:

$$\sum_{i=1}^n \alpha_i \exp \left[\lambda_i \int_{\pi_0}^{\mu_i} d \log s \right] (\pi_0 - \mu_i) = 0, \text{ or, equivalently:} \quad (38)$$

$$\sum_{i=1}^n \alpha_i \left(\frac{\mu_i}{\pi_0} \right)^{\lambda_i} (\pi_0 - \mu_i) = 0 \quad (39)$$

Note that Equation 39 is identical in form to Equation 29, but with the relevant arc elasticities substituted for point elasticities.

The general result, as stated below, then follows directly from Result 5.2:

Result 5.3. *Suppose there are n risk groups with risks $\mu_1 < \mu_2 < \dots < \mu_n$ with demand elasticities $\epsilon_1(\pi), \epsilon_2(\pi), \dots, \epsilon_n(\pi)$, such that $\lambda_1, \lambda_2, \dots, \lambda_n$ are the respective arc elasticities under pooled equilibrium. Define $\lambda_{lo} = \max_{v \leq 1} f(v)$ and $\lambda_{hi} = \min_{v > 1} f(v)$. Then $\lambda_{lo} < 1$ and $\lambda_{hi} \geq \lambda_{lo} \Rightarrow C \geq 1$.*

In other words, under pooled equilibrium, as long as the arc elasticities of lower risk-groups (paying more than their fair actuarial premium) do not exceed 1 and the arc elasticities of the higher risk-groups (paying less than their fair actuarial premium) in all cases exceed those of the lower risk-groups, then loss coverage under pooling is higher than under full risk classification.

For the special case of $\epsilon_i(\pi)$ being non-decreasing functions of premium π and bounded above by 1, where $\epsilon_1(\pi) \leq \epsilon_2(\pi) \leq \dots \leq \epsilon_n(\pi)$ the required conditions for Result 5.3 are automatically satisfied. This is because the arc elasticities of demand, being an *average* of the underlying point elasticities, satisfy: $0 < \lambda_1 \leq \lambda_2 \leq \dots \leq \lambda_n < 1$, which then implies that $\lambda_{lo} < 1$ and $\lambda_{hi} \geq \lambda_{lo}$ and thus $C \geq 1$.

6. DISCUSSION

6.1 EMPIRICAL ESTIMATES

The results obtained in Section 5 suggest that loss coverage will be higher under pooling than under full risk classification, if

1. elasticity (or arc elasticity, if elasticity is not constant) is less than 1 for all lower risk-groups; and
2. elasticity (or arc elasticity, if elasticity is not constant) for all higher risk-groups exceeds that for all lower risk-groups,

where arc elasticities are logarithmic averages of demand elasticities over the arc, from true risk price to the equilibrium pooled price.

Are these conditions likely to be satisfied in the real world?

For the first condition, Table 1 shows some relevant empirical estimates for insurance demand elasticities.⁴ It can be seen that most estimates are of magnitude significantly less than 1. Whilst the various contexts in which these estimates were made may not correspond closely to the set-up in this paper, the figures are at least suggestive of the possibility that the first condition may often be satisfied.

Table 1: Estimates of demand elasticity for various insurance markets

Market & country	Demand elasticities	Authors
Term life insurance, USA	-0.66	Viswanathan et al. (2006)
Yearly renewable term life, USA	-0.4 to -0.5	Pauly et al. (2003)
Whole life insurance, USA	-0.71 to -0.92	Babbel (1985)
Health insurance, USA	0 to -0.2	Chernew et al. (1997), Blumberg et al. (2001), Buchmueller and Ohri (2006)
Health insurance, Australia	-0.35 to -0.50	Butler (1999)
Farm crop insurance, USA	-0.32 to -0.73	Goodwin (1993)

For the second condition, we know of no empirical evidence that insurance demand elasticities are higher (or lower) for higher risks. However, as noted earlier in the paper, this condition may be plausible in that it is consistent with the income effect on demand.

⁴Demand elasticity is defined as a positive constant in this paper for convenience, but estimates in empirical papers are generally given with the negative sign, so Table 1 quotes them in that form.

6.2 APPLICABILITY OF RESULTS

In the derivation of proportional demand from heterogeneous individual preferences in Section 3, we made assumptions that all individuals had the same initial wealth W and the same potential loss amount L . It is worth noting that the the results about demand functions in Section 5 apply more generally. In particular, *if the demand function is given*, the results in Section 5 require no assumptions whatsoever about individuals' initial wealth W . The form of the demand functions specified in Section 5 does however require that all individuals have the same potential loss amount L , and that any insurance fully covers this loss amount.

7. CONCLUSIONS

Loss coverage is defined as the expected population losses compensated by insurance at market equilibrium. We suggest that if the social purpose of insurance is to compensate the population's losses, loss coverage may be an intuitively appealing metric for evaluation of different risk classification schemes.

When risk classification is banned, so that insurers have to pool all risks at a single price, this can lead to adverse selection. Adverse selection is associated with a fall in the number of insured individuals compared with that obtained under full risk classification. However, adverse selection is also associated with a shift in coverage towards higher risks. If the shift is large enough, it can more than offset the fall in numbers insured, so that loss coverage is increased. We suggest that from a social perspective, this possibility might be seen as a good outcome from adverse selection.

Whether loss coverage is in fact increased when all risks are pooled depends on the response of higher and lower risks to changes in the prices they face, that is the demand elasticities of higher and lower risks.

This paper has stated demand elasticity conditions which ensure that loss coverage will be higher under pooling than under fully risk-differentiated premiums. The conditions were stated for successively more general contexts: first, for iso-elastic demand with a single elasticity parameter; second, for iso-elastic demand with different elasticity parameters for different risk-groups; and third, for any downward-sloping demand functions (including different ones for different risk-groups). The demand elasticities required for loss coverage to be higher under pooling seem realistic for some insurance markets.

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APPENDICES

A. MICROFOUNDATION OF INSURANCE DEMAND

Suppose all individuals have the same initial wealth W and the same potential loss amount L (with $L \leq W$), and risk preferences driven by a power utility function:

$$U_\gamma(w) = \left[\frac{w - (W - L)}{L} \right]^\gamma, \quad (40)$$

so that $U_\gamma(W) = 1$ and $U_\gamma(W - L) = 0$. This particular form of utility function leads to:

$$\text{relative risk aversion coefficient: } -w \frac{U_\gamma''(w)}{U_\gamma'(w)} = (1 - \gamma) \left[\frac{w}{w - (W - L)} \right]. \quad (41)$$

Given that all individuals have the same W and same L , Equation 41 implies that the heterogeneity in preferences between individuals can be captured by the randomness of γ , the risk aversion parameter.⁵ As outlined in Section 3, we define a positive random variable Γ , and individual risk preferences γ are then instances drawn from the distribution of Γ .

Inserting the utility function in Equation 40 into Equation 7, the demand for insurance at a given premium π is then:

$$d(\pi) = \text{P} [U_\Gamma(W - \pi L) > 1 - \mu], \quad (42)$$

$$= \text{P} [(1 - \pi)^\Gamma > 1 - \mu], \quad (43)$$

$$= \text{P} [\Gamma \log(1 - \pi) > \log(1 - \mu)], \text{ as log is monotonic,} \quad (44)$$

$$= \text{P} \left[\Gamma < \frac{\log(1 - \mu)}{\log(1 - \pi)} \right], \text{ as } \log(1 - \pi) < 0. \quad (45)$$

So, given an observed (proportional) insurance demand function, which is non-increasing in premium, π , the underlying random variable Γ has the following distribution function:

$$F_\Gamma(\gamma) = \text{P} [\Gamma < \gamma] = d(1 - (1 - \mu)^{1/\gamma}). \quad (46)$$

Since $1 - (1 - \mu)^{1/\gamma} \approx \mu/\gamma$ for small μ , then for small premiums we can approximate equation (46) by the simpler:

$$F_\Gamma(\gamma) = \text{P} [\Gamma < \gamma] = d \left(\frac{\mu}{\gamma} \right). \quad (47)$$

Note that $F_\Gamma(\gamma)$ is a non-decreasing function and lies between 0 and 1.

Of course, for F_Γ to be a valid distribution function, we would also require $\lim_{\gamma \rightarrow 0} F_\Gamma(\gamma) = 0$ and $\lim_{\gamma \rightarrow \infty} F_\Gamma(\gamma) = 1$, or equivalently, $\lim_{\pi \rightarrow \infty} d(\pi) = 0$ and $\lim_{\pi \rightarrow 0} d(\pi) = 1$, which appear to

⁵In the special case $W = L$, Equation 41 reduces to $(1 - \gamma)$, the familiar case of constant relative risk aversion.

be reasonable assumptions. However, empirical observations are unlikely to be available for these extreme cases, so it is only possible to model insurance purchasing behaviour over the range of premiums observed in the market, with appropriate extrapolations at the limiting extremes.

We formally present this result as follows:

Result A.1. *Given an observed proportional insurance demand, $d(\pi)$, which is a valid probability and non-increasing in π , heterogeneity of risk preferences driven by the power utility function and characterised by the random parameter Γ with the distribution function given by Equation 47 produces the observed demand for insurance for small premiums.*

We illustrate the above result using iso-elastic demand: Suppose Γ has the following distribution:

$$F_{\Gamma}(\gamma) = \text{P}[\Gamma \leq \gamma] = \begin{cases} 0 & \text{if } \gamma < 0 \\ \tau \gamma^{\lambda} & \text{if } 0 \leq \gamma \leq (1/\tau)^{1/\lambda} \\ 1 & \text{if } \gamma > (1/\tau)^{1/\lambda}, \end{cases} \quad (48)$$

where τ and λ are positive parameters. Note that $\tau = \lambda = 1$ leads to a uniform distribution. λ controls the shape of the distribution function and τ controls the range over which Γ takes its values. Then the demand for insurance, as given in Equation (47), takes the form:

$$d(\pi) = \tau \left(\frac{\mu}{\pi} \right)^{\lambda}, \quad (\text{subject to a cap of } 1) \quad (49)$$

which corresponds to iso-elastic demand, the constant demand elasticity being:

$$\epsilon(\pi) = -\frac{\partial \log(d(\pi))}{\partial \log \pi} = \lambda. \quad (50)$$

The parameter τ can also be interpreted as the *fair-premium demand*, that is the demand when an actuarially fair premium is charged.

B. EQUAL ISO-ELASTIC DEMAND ELASTICITIES

Using the formulation of Section 5.1, we prove the following result.

Result B.1. *Let V be a positive random variable and λ be a positive constant, such that $\mathbf{E}[V^\lambda] = \mathbf{E}[V^{\lambda+1}]$. Then:*

$$\lambda \lesseqgtr 1 \Rightarrow \mathbf{E}[V^\lambda] \gtrless \mathbf{E}[V]. \quad (51)$$

Proof. Case: $\lambda = 1$: It follows directly from the definition.

Case: $0 < \lambda < 1$: Holder's inequality states that, if $1 < p, q < \infty$ where $1/p + 1/q = 1$, for positive random variables X, Y with $\mathbf{E}[X^p], \mathbf{E}[Y^q] < \infty$:

$$(\mathbf{E}[X^p])^{1/p} (\mathbf{E}[Y^q])^{1/q} \geq \mathbf{E}[XY]. \quad (52)$$

Setting $1/p = \lambda$, $1/q = 1 - \lambda$, $X = V^{\lambda^2}$ and $Y = V^{1-\lambda^2}$, Holder's inequality gives:

$$\left(\mathbf{E}\left[V^{\lambda^2 \frac{1}{\lambda}}\right]\right)^\lambda \left(\mathbf{E}\left[V^{(1-\lambda^2) \frac{1}{1-\lambda}}\right]\right)^{1-\lambda} \geq \mathbf{E}\left[V^{\lambda^2} V^{1-\lambda^2}\right], \quad (53)$$

$$\Rightarrow \left(\mathbf{E}\left[V^\lambda\right]\right)^\lambda \left(\mathbf{E}\left[V^{\lambda+1}\right]\right)^{1-\lambda} \geq \mathbf{E}[V], \quad (54)$$

$$\Rightarrow \mathbf{E}\left[V^\lambda\right] \geq \mathbf{E}[V], \quad \text{since } \mathbf{E}\left[V^\lambda\right] = \mathbf{E}\left[V^{\lambda+1}\right]. \quad (55)$$

Case: $\lambda > 1$: Young's inequality states that, for $a, b \geq 0$ and $p, q > 0$ such that $1/p + 1/q = 1$:

$$ab \leq \frac{a^p}{p} + \frac{b^q}{q}. \quad (56)$$

Setting $p = \lambda$, $q = \frac{\lambda}{\lambda-1}$, $a = V^{\frac{1}{\lambda}}$ and $b = V^{\lambda - \frac{1}{\lambda}}$, Young's inequality gives:

$$V^{\frac{1}{\lambda}} V^{\lambda - \frac{1}{\lambda}} \leq \frac{1}{\lambda} V^{\frac{1}{\lambda} \lambda} + \frac{\lambda-1}{\lambda} V^{(\lambda - \frac{1}{\lambda}) \frac{\lambda}{\lambda-1}}, \quad (57)$$

$$\Rightarrow V^\lambda \leq \frac{1}{\lambda} V + \frac{\lambda-1}{\lambda} V^{\lambda+1}, \quad (58)$$

$$\Rightarrow \mathbf{E}\left[V^\lambda\right] \leq \frac{1}{\lambda} \mathbf{E}[V] + \frac{\lambda-1}{\lambda} \mathbf{E}\left[V^{\lambda+1}\right], \quad (59)$$

$$\Rightarrow \mathbf{E}\left[V^\lambda\right] \leq \mathbf{E}[V], \quad \text{since } \mathbf{E}\left[V^\lambda\right] = \mathbf{E}\left[V^{\lambda+1}\right]. \quad (60)$$

□

Result 5.1 follows directly from Result B.1 by noting that: $C = \mathbf{E}[V^\lambda] / \mathbf{E}[V]$.

C. DIFFERENT ISO-ELASTIC DEMAND ELASTICITIES

Using the formulation of Section 5.2, we prove the following result:

Result C.1. *Let V be a positive random variable and $f(v)$ be a positive function, such that $E[V^{f(V)}] = E[V^{f(V)+1}]$. Define $\lambda_{lo} = \max_{v \leq 1} f(v)$ and $\lambda_{hi} = \min_{v > 1} f(v)$. Then:*

$$\lambda_{lo} < 1 \text{ and } \lambda_{hi} \geq \lambda_{lo} \Rightarrow E[V^{f(V)}] \geq E[V]. \quad (61)$$

Proof. Holder's inequality states that, if $1 < p, q < \infty$ where $1/p + 1/q = 1$, for positive random variables X, Y with $E[X^p], E[Y^q] < \infty$:

$$(E[X^p])^{1/p} (E[Y^q])^{1/q} \geq E[XY]. \quad (62)$$

For any λ , such that $0 < \lambda < 1$, set $1/p = \lambda$, $1/q = 1 - \lambda$, $X = V^{\lambda f(V)}$ and $Y = V^{(1-\lambda)(f(V)+1)}$, Holder's inequality gives:

$$\left(E[V^{f(V)}]\right)^\lambda \left(E[V^{f(V)+1}]\right)^{1-\lambda} \geq E[V^{\lambda f(V)} V^{(1-\lambda)(f(V)+1)}], \quad (63)$$

$$\Rightarrow E[V^{f(V)}] \geq E[V^{f(V)+1-\lambda}], \quad \text{since } E[V^{f(V)}] = E[V^{f(V)+1}]. \quad (64)$$

The relationship in Equation 64 holds for any positive $\lambda < 1$. Now, set $\lambda = \lambda_{lo} < 1$.

Case $V < 1$:

$$\lambda_{lo} = \max_{v \leq 1} f(v) \Rightarrow f(V) \leq \lambda_{lo} = \lambda \Rightarrow f(V) + 1 - \lambda \leq 1 \Rightarrow V^{f(V)+1-\lambda} \geq V. \quad (65)$$

Case $V = 1$:

$$V^{f(V)+1-\lambda} = V. \quad (66)$$

Case $V > 1$:

$$\lambda_{hi} = \min_{v > 1} f(v) \Rightarrow f(V) \geq \lambda_{hi} \geq \lambda_{lo} = \lambda \Rightarrow f(V) + 1 - \lambda \geq 1 \Rightarrow V^{f(V)+1-\lambda} \geq V. \quad (67)$$

Hence, $V^{f(V)+1-\lambda} \geq V$ for all cases, which implies

$$E[V^{f(V)+1-\lambda}] \geq E[V]. \quad (68)$$

Combining Equations 64 and 68, we have:

$$E[V^{f(V)}] \geq E[V]. \quad (69)$$

□

Result 5.2 follows directly from Result C.1 by noting that the loss coverage ratio in this case is: $C = E[V^{f(V)}] / E[V]$.